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## Odd Generalized N-H Generated Family of Distributions with Application to Exponential Model



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### Abstract

A new family of distributions called the odd generalized N-H is introduced and studied. Four new special models are presented. Some mathematical properties of the odd generalized N-H family are studied. Explicit expressions for the moments, probability weighted, quantile function, mean deviation, order statistics and Rényi entropy are investigated. Characterizations based on the truncated moments, hazard function and conditional expectations are presented for the generated family. Parameter estimates of the family are obtained based on maximum likelihood procedure. Two real data sets are employed to show the usefulness of the new family.

**Key Words:** N-H distribution; Order statistics; Maximum likelihood method.

### 1. Introduction

In recent years, several classes have been defined by adding one or more parameters to generate new distributions. These distributions extend well-known distributions as well as provide great flexibility to model specific real data. Some of the well-known generators are the beta-G by Eugene et al. (2002), Kumaraswamy-G by Cordeiro and de Castro (2011), exponentiated generalized-G by Cordeiro et al. (2013), Transformed-Transformer (T-X) by Alzaatreh et al. (2013), Weibull-G by Bourguignon et al. (2014), exponentiated half-logistic-G by Cordeiro et al. (2014), odd generalized exponential by Tahir (2015), beta odd log-logistic generalized by Cordeiro et al. (2016), Garhy-G by Elgarhy et al. (2016), exponentiated Weibull-G by Hassan and Elgarhy (2016a), Kumaraswamy Weibull-G by Hassan and Elgarhy (2016b), exponentiated extended-G by Elgarhy et al. (2017), type II half logistic-G by Hassan et al. (2017), odd Burr generalized by Alizadeh et al. (2017), generalized odd log-logistic by Cordeiro et al. (2017), a new generalized odd log-logistic by Haghbin et al. (2017), odd Lindley-G by Gomes-Silva (2017), among others.

In this paper, we introduce a new generated family of distributions using the NH distribution as a generator. The NH distribution is introduced by (Nadarajah and Haghighi, 2011) which has the following cumulative distribution function (cdf)

$$G(t) = 1 - e^{\{1-(1+\lambda t)^\alpha\}} = 1 - Ce^{-(1+\lambda t)^\alpha}, t > 0, \lambda, \alpha > 0, \quad (1)$$

where  $C = e$ . The associated probability density function (pdf) corresponding to (1) is as follows

$$g(t) = \alpha\lambda C(1 + \lambda t)^{\alpha-1}e^{-(1+\lambda t)^\alpha}, \tag{2}$$

where  $\lambda$  is the scale parameter and  $\alpha$  is the shape parameter. We are interested in modeling the random variable  $X$  of this odds using the NH model (with scale parameter  $\lambda = 1$ ) given by (1). The cdf of OGNH-G family can be expressed as follows

$$F(x) = \int_0^{\left(\frac{G(x;\zeta)}{1-G(x;\zeta)}\right)^\beta} C\alpha(1+t)^{\alpha-1}e^{-(1+t)^\alpha} dt = 1 - Ce^{-\left(1+\left(\frac{G(x;\zeta)}{1-G(x;\zeta)}\right)^\beta\right)^\alpha}, \quad x \in \mathbb{R}, \alpha, \beta > 0, \tag{3}$$

where,  $\alpha, \beta$  are two shape parameters, and  $G(x; \zeta)$  is a baseline cdf, which depends on a parameter vector  $\zeta$ . The distribution function (3) provides a broadly odd half logistic generated distributions. Therefore, the pdf of the OGNH-G family is as follows

$$f(x) = \alpha\beta Cg(x; \zeta) \frac{G(x;\zeta)^{\beta-1}}{(1-G(x;\zeta))^{\beta+1}} \left(1 + \left(\frac{G(x;\zeta)}{1-G(x;\zeta)}\right)^\beta\right)^{\alpha-1} e^{-\left(1+\left(\frac{G(x;\zeta)}{1-G(x;\zeta)}\right)^\beta\right)^\alpha}. \tag{4}$$

Hereafter, a random variable  $X$  has pdf (4) will be denoted by  $X \sim \text{OGNH} - G$ .

The survival function, hazard rate and reversed hazard rate functions are, respectively, given by

$$\begin{aligned} \bar{F}(x) &= Ce^{-\left(1+\left(\frac{G(x;\zeta)}{1-G(x;\zeta)}\right)^\beta\right)^\alpha}, \\ h(x) &= \alpha\beta g(x; \zeta) \frac{G(x;\zeta)^{\beta-1}}{(1-G(x;\zeta))^{\beta+1}} \left(1 + \left(\frac{G(x;\zeta)}{1-G(x;\zeta)}\right)^\beta\right)^{\alpha-1}, \end{aligned}$$

and

$$\tau(x) = \frac{\alpha\beta Cg(x; \zeta) \frac{G(x;\zeta)^{\beta-1}}{(1-G(x;\zeta))^{\beta+1}} \left(1 + \left(\frac{G(x;\zeta)}{1-G(x;\zeta)}\right)^\beta\right)^{\alpha-1} e^{-\left(1+\left(\frac{G(x;\zeta)}{1-G(x;\zeta)}\right)^\beta\right)^\alpha}}{1 - Ce^{-\left(1+\left(\frac{G(x;\zeta)}{1-G(x;\zeta)}\right)^\beta\right)^\alpha}}.$$

This paper can be sorted as follows. In the next section, Characterizations based on the truncated moments, hazard function and conditional expectations are presented for the OGNH-G family. Section 3 provides some general mathematical properties of the family. Section 4 gives the estimation of the parameters of the family using the maximum likelihood method. In Section 5, some new special models of the generated family are considered. Some statistical properties, estimation and simulation study for odd generalized NH exponential model are derived in Section 6. The application of the odd generalized NH exponential distribution to a real data set is presented in Section 7. At the end, concluding remarks are outlined in Section 8.

## 2. Characterizations of OGNH-G Distribution

This section is devoted to the characterizations of the OGNH-G distribution in the following directions: (i) based on the ratio of two truncated moments and (ii) in terms of the hazard. Note that (i) can be employed also when the cdf does not have a closed form. We also like to point out that the characterization (i) is stable in the sense of weak convergence. We present our characterizations (i)-(ii) in two subsections.

### 2.1. Characterizations based on two truncated moments

This subsection is devoted to the characterizations of OGNH-G distribution based on the ratio of two truncated moments. Our first characterization employs a theorem due to Glänzel (1987), see Theorem 1 of Appendix A. The result, however, holds also when the interval  $H$  is not closed, since the condition of the Theorem is on the interior of  $H$ .

**Proposition 2.1.** Let  $X: \Omega \rightarrow R$  be a continuous random variable and let

$$q_1(x) = \tilde{G}(x; \xi)^{\alpha+1} \left(1 + \left(\frac{G(x; \xi)}{\tilde{G}(x; \xi)}\right)^\beta\right)^{1-\alpha} \exp\left\{\left(1 + \left(\frac{G(x; \xi)}{\tilde{G}(x; \xi)}\right)^\beta\right)^\alpha\right\} \text{ and } q_2(x) = q_1(x)G(x; \xi)^\alpha \text{ for } X \in \mathbb{R}.$$

The random variable  $X$  has pdf (4) if and only if the function  $\xi$  defined in Theorem 1 is of the form

$$\xi(x) = \frac{1}{2} [1 + G(x; \xi)^\alpha], \quad X \in \mathbb{R}.$$

**Proof.** Suppose the random variable  $X$  has pdf (4), then

$$(1 - F(x))E[q_1(x)/X \geq x] = C\beta[1 - G(x; \xi)^\alpha], \quad X \in \mathbb{R},$$

and

$$(1 - F(x))E[q_2(x)/X \geq x] = \frac{C\beta}{2} [1 - G(x; \xi)^{2\alpha}], \quad X \in \mathbb{R}.$$

Further,

$$\xi(x)q_1(x) - q_2(x) = \frac{q_1(x)}{2} [1 - G(x; \xi)^\alpha] > 0, \quad \text{for } X \in \mathbb{R}.$$

Conversely, if  $\xi$  is of the above form, then

$$S'(x) = \frac{\xi'(x)q_1(x)}{\xi(x)q_1(x) - q_2(x)} = \frac{\alpha g(x; \xi)G(x; \xi)^{\alpha-1}}{1 - G(x; \xi)^\alpha}, \quad X \in \mathbb{R},$$

and consequently

$$S(x) = -\log[1 - G(x; \xi)^\alpha], \quad X \in \mathbb{R}.$$

Now, according to Theorem 1,  $X$  has density (4).

**Corollary 2.1.** Let  $X: \Omega \rightarrow R$  be a continuous random variable and let  $q_1(x)$  be as in Proposition A.1. The random variable  $X$  has pdf (4) if and only if there exist functions  $q_2(x)$  and  $\xi$  defined in Theorem 1 satisfying the following differential equation

$$\frac{\xi'(x)q_1(x)}{\xi(x)q_1(x) - q_2(x)} = \frac{\alpha g(x; \xi)G(x; \xi)^{\alpha-1}}{1 - G(x; \xi)^\alpha}, \quad X \in \mathbb{R}.$$

**Corollary 2.2.** The general solution of the differential equation in Corollary A.1 is

$$\xi(x) = (1 - G(x; \phi)^\alpha)^{-1} \left[ - \int \alpha g(x; \xi)G(x; \phi)^{\alpha-1} (q_1(x))^{-1} q_2(x) dx + D \right],$$

where  $D$  is a constant. We like to point out that one set of functions satisfying the above differential equation is given in Proposition A.1 with  $D=1/2$ . Clearly, there are other triplets  $(q_1, q_2, \xi)$  which satisfy conditions of Theorem 1.

## 2.2. Characterization in terms of hazard function

The hazard function,  $h_F$ , of a twice differentiable distribution function,  $F$ , satisfies the following first order differential equation

$$\frac{f'(x)}{f(x)} = \frac{h'_F(x)}{h_F(x)} - h_F(x).$$

It should be mentioned that for many univariate continuous distributions, the above equation is the only differential equation available in terms of the hazard function. In this subsection we present a non-trivial characterization of OGNH-G distribution in terms of the hazard function.

**Proposition 2.2.1** Let  $X: \Omega \rightarrow \mathbb{R}$  be a continuous random variable. The random variable  $X$  has pdf (4) if and only if its hazard function  $h_F$ , satisfies the following differential equation

$$h'_F(x) - \frac{g'(x; \xi)}{g(x; \xi)} h_F(x) = \alpha\beta g(x; \xi) \frac{d}{dx} \left\{ \frac{[G(x; \xi)\{\bar{G}(x; \xi)^\beta + G(x; \xi)^\beta\}]^{\alpha-1}}{\bar{G}(x; \xi)^{\alpha\beta + \alpha - \beta + 1}} \right\}, \quad X \in \mathbb{R}.$$

**Proof.** If  $X$  has pdf (4), then clearly the above differential equation holds. If the differential equation holds, then

$$\frac{d}{dx} \{g(x; \xi)^{-1} h_F(x)\} = \alpha\beta \frac{d}{dx} \left\{ \frac{G(x; \xi)^{\alpha-1}}{\bar{G}(x; \xi)^{\alpha+1}} \left( 1 + \left( \frac{G(x; \xi)}{\bar{G}(x; \xi)} \right)^\beta \right)^{\alpha-1} \right\},$$

from which we arrive at the hazard function corresponding to the pdf (4).

**Remark 2.1.** For  $\beta = 1$ , we have a much simpler differential equation

$$h'_F(x; \xi) - \frac{g'(x; \xi)}{g(x; \xi)} h_F(x; \xi) = \alpha(\alpha + 1)g(x; \xi)^2 \bar{G}(x; \xi)^{-\alpha-2}, \quad X \in \mathbb{R}.$$

### 3. Some Statistical Properties

This section provides some statistical properties of OGNH-G family of distributions.

#### 3.1. Quantile function

Let  $X$  denotes a random variable has the pdf (4), the quantile function, say  $Q(u)$  of  $X$  is given by

$$Q(u) = G^{-1} \left\{ \frac{[1 - [1 - \ln[1 - u]]^{\frac{1}{\alpha}}]^{\frac{1}{\beta}}}{1 + [1 - [1 - \ln[1 - u]]^{\frac{1}{\alpha}}]^{\frac{1}{\beta}}} \right\}, \tag{5}$$

where,  $u$  is a uniform distribution on the interval (0,1) and  $G^{-1}(\cdot)$  is the inverse function of  $G(\cdot)$ .

#### 3.2. Useful representation

In this subsection, a useful expansion of the pdf and cdf for OGNH-G is provided.

Since the exponential series is

$$e^{-\left(1 + \left(\frac{G(x; \zeta)}{1 - G(x; \zeta)}\right)^\beta\right)^\alpha} = \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} \left(1 + \left(\frac{G(x; \zeta)}{1 - G(x; \zeta)}\right)^\beta\right)^{\alpha i}. \tag{6}$$

Inserting (6) in (4) then,

$$f(x) = \alpha\beta C g(x; \zeta) \frac{G(x; \zeta)^{\beta-1}}{(1 - G(x; \zeta))^{\beta+1}} \sum_{i=0}^{\infty} \frac{(-1)^i}{i!} \left(1 + \left(\frac{G(x; \zeta)}{1 - G(x; \zeta)}\right)^\beta\right)^{\alpha(i+1)-1}.$$

By using binomial theory

$$(1 + z)^\beta = \sum_{j=0}^{\infty} \binom{\beta}{j} z^j, \tag{7}$$

for  $|z| < 1$ , and  $\beta$  is a positive real non integer. Then, by applying the binomial theorem (7) in the previous density function of OGNH-G family becomes

$$f(x) = \alpha\beta C g(x; \zeta) \sum_{i,j=0}^{\infty} \frac{(-1)^i}{i!} \binom{\alpha(i+1)-1}{j} \frac{G(x; \zeta)^{\beta(j+1)-1}}{(1-G(x; \zeta))^{\beta(j+1)+1}}. \tag{8}$$

Since, the binomial expansion

$$(1-z)^{-\beta} = \sum_{k=0}^{\infty} \binom{\beta+k-1}{k} z^k, \tag{9}$$

By inserting (9) in (8) then, the pdf (8) can be written as follows

$$f(x) = \sum_{k=0}^{\infty} \eta_k g(x; \zeta) G(x; \zeta)^{\beta(j+1)+k-1}, \tag{10}$$

where,

$$\eta_k = \sum_{i,j=0}^{\infty} \frac{\alpha\beta C (-1)^i}{i!} \binom{i}{j} \binom{\alpha(i+1)-1}{j} \binom{\beta(j+1)+k}{k}.$$

Another formula can be extracted from pdf (10), which gives the following infinite linear combination

$$f(x) = \sum_{k=0}^{\infty} W_k h_{\beta(j+1)+k}(x), \tag{11}$$

where,  $W_k = \frac{\eta_k}{(\beta(j+1)+k)}$ , and  $h_a(x) = ag(x; \zeta)G(x; \zeta)^{a-1}$ , is the exponentiated-generated (exp-G) density with power parameter  $a$ . Further, an expansion for the  $[F(x)]^h$  is derived, for  $h$  is integer, again, the binomial and exponential expansions is worked out.

$$[F(x)]^h = \sum_{z=0}^{\infty} s_z G(x; \zeta)^{\beta u+z}, \tag{12}$$

where,

$$s_z = \sum_{i=0}^h \sum_{j,u=0}^{\infty} \frac{i^j C^i (-1)^{i+j}}{j!} \binom{h}{i} \binom{\alpha j}{u} \binom{\beta u+z-1}{z}.$$

### 3.3. The probability weighted moments

For a random variable  $X$ , the probability-weighted moments (PWMs), denoted by  $\tau_{r,s}$ , can be calculated through the following relation

$$\tau_{r,s} = E[X^r F(x)^s] = \int_{-\infty}^{\infty} x^r f(x) (F(x))^s dx. \tag{13}$$

The PWMs of OGNH-G is obtained by substituting (10) and (12) into (13), and replacing  $h$  with  $s$ , as follows

$$\tau_{r,s} = \int_{-\infty}^{\infty} \sum_{k,z=0}^{\infty} s_z \eta_k x^r g(x; \zeta) (G(x; \zeta))^{\beta(j+u+1)+k+z-1} dx.$$

Then,

$$\tau_{r,s} = \sum_{k,z=0}^{\infty} s_z \eta_k \tau_{r,\beta(j+u+1)+k+z-1},$$

where,  $\tau_{r,\beta(j+u+1)+k+z-1} = \int_{-\infty}^{\infty} x^r g(x; \zeta) (G(x; \zeta))^{\beta(j+u+1)+k+z-1} dx$ .

### 3.4. Moments

Since the moments are necessary and important in any statistical analysis, especially in applications. Therefore, we derive the  $r$ th moment for the OGNH-G family. If  $X$  has the pdf (10), then  $r$ th moment is obtained as follows

$$\mu'_r = \int_{-\infty}^{\infty} x^r f(x) dx = \int_{-\infty}^{\infty} \sum_{k=0}^{\infty} \eta_k x^r g(x; \zeta) G(x; \zeta)^{\beta(j+1)+k-1} dx.$$

Then,

$$\mu'_r = \sum_{k=0}^{\infty} \eta_k \tau_{r,\beta(j+1)+k-1}.$$

For a random variable  $X$ , it is known that, the moment generating function is defined as

$$M_X(t) = \sum_{r=0}^{\infty} \frac{t^r}{r!} \mu_r = \sum_{r,k=0}^{\infty} \frac{t^r}{r!} \eta_k \tau_{r,\beta(j+1)+k-1}.$$

**3.5. The mean deviation**

For random variable  $X$  with pdf  $f(x)$ , cdf  $F(x)$ , the mean deviation about the mean and mean deviation about the median, are defined by

$$\delta_1 = 2\mu F(\mu) - 2T(\mu) \text{ and } \delta_2 = \mu - 2T(M),$$

where,  $\mu = E(X)$ ,  $M = \text{Median}(X)$ , and  $T(q) = \int_{-\infty}^q xf(x)dx$  which is the first incomplete moment.

**3.6. Order statistics**

Order statistics have been extensively applied in many fields of statistics, such as reliability and life testing. Let  $X_1, X_2, \dots, X_n$  be independent and identically distributed (i.i.d) random variables with their corresponding continuous distribution function  $F(x)$ . Let  $X_{1:n} < X_{2:n} < \dots < X_{n:n}$  the corresponding ordered random sample from a population of size  $n$ . According to (David, 1981), the pdf of the  $s$ th order statistic, is defined as

$$f_{s:n}(x) = \frac{f(x)}{B(k,n-k+1)} \sum_{v=0}^{n-s} (-1)^v \binom{n-s}{v} F(x)^{v+s-1}, \tag{14}$$

where,  $B(\cdot, \cdot)$  stands for beta function. The pdf of the  $s$ th order statistic for OGNH-G family is derived by substituting (10) and (12) in (14), replacing  $h$  with  $v + k - 1$ ,

$$f_{s:n}(x) = \frac{g(x;\zeta)}{B(s,n-s+1)} \sum_{v=0}^{n-s} \sum_{k,z=0}^{\infty} \eta_k p_{z,v} G(x;\zeta)^{\beta(j+u+1)+k+z-1}, \tag{15}$$

where  $p_{z,v} = (-1)^v \binom{n-s}{v} s_z$ ,  $g(\cdot)$  and  $G(\cdot)$  are the pdf and cdf of the OGNH-G family, respectively. Further, the  $r$ th moment of  $s$ th order statistics for OGNH-G is defined family by:

$$E(X_{s:n}^r) = \int_{-\infty}^{\infty} x^r f_{s:n}(x) dx. \tag{16}$$

By substituting (15) in (16), leads to

$$E(X_{s:n}^r) = \frac{1}{B(s,n-s+1)} \sum_{v=0}^{n-s} \sum_{k,z=0}^{\infty} \eta_k p_{z,v} \int_{-\infty}^{\infty} x^r g(x;\zeta) G(x;\zeta)^{\beta(j+u+1)+k+z-1} dx.$$

Then,

$$E(X_{s:n}^r) = \frac{1}{B(s,n-s+1)} \sum_{v=0}^{n-s} \sum_{k,z=0}^{\infty} \eta_k p_{z,v} \tau_{r,\beta(j+u+1)+k+z-1}.$$

**3.7 Rényi entropy**

The entropy of a random variable  $X$  is a measure of variation of uncertainty and has been used in many fields such as physics, engineering and economics. As mentioned by (Rényi 1961), the Rényi entropy is defined by

$$I_{\delta}(X) = \frac{1}{1-\delta} \log \int_{-\infty}^{\infty} f(x)^{\delta} dx, \delta > 0 \text{ and } \delta \neq 1.$$

By applying the binomial and exponential theory in the pdf (4), then the pdf  $f(x)^{\delta}$  can be expressed as follows

$$f(x)^{\delta} = \sum_{k=0}^{\infty} t_k g(x;\zeta)^{\delta} G(x;\zeta)^{j+k},$$

where

$$t_k = \sum_{i,j=0}^{\infty} \frac{(\alpha\beta C)^{\delta} (-\delta)^i}{i!} \binom{\alpha(i+\delta)-\delta}{j} \binom{\beta(j+\delta)+\delta+k-1}{k}.$$

Therefore, the Rényi entropy of OGNH generated family of distributions is given by

$$I_{\delta}(X) = \frac{1}{1-\delta} \log \sum_{k=0}^{\infty} t_k \int_{-\infty}^{\infty} g(x;\zeta)^{\delta} G(x;\zeta)^{\beta(j+\delta)+k-\delta} dx.$$

#### 4. Maximum Likelihood Method

This section deals with the maximum likelihood estimators of the unknown parameters for the OGNH-G family of distributions on the basis of complete samples. Let  $X_1, \dots, X_n$  be the observed values from the ONH-G family with set of parameter  $\Phi = (\alpha, \beta, \zeta)^T$ . The log-likelihood function for parameter vector  $\Phi = (\alpha, \beta, \zeta)^T$  is obtained as follows

$$\begin{aligned} \ln L(\Phi) &= n \ln \alpha + n \ln \beta + \sum_{i=1}^n \ln[g(x_i; \zeta)] + (\beta - 1) \sum_{i=1}^n \ln[G(x_i; \zeta)] - (\beta + 1) \sum_{i=1}^n \ln[1 - G(x_i; \zeta)] \\ &+ (\alpha - 1) \sum_{i=1}^n \ln \left[ 1 + \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)^\beta \right] + \sum_{i=1}^n \left\{ 1 - \left( 1 + \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)^\beta \right)^\alpha \right\}. \end{aligned}$$

The elements of the score function  $U(\Phi) = (U_\alpha, U_\beta, U_{\zeta_k})$  are given by

$$\begin{aligned} U_\alpha &= \frac{n}{\alpha} + \sum_{i=1}^n \ln \left[ 1 + \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)^\beta \right] - \sum_{i=1}^n \left( 1 + \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)^\beta \right)^\alpha \ln \left( 1 + \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)^\beta \right), \\ U_\beta &= \frac{n}{\beta} + \sum_{i=1}^n \ln[G(x_i; \zeta)] - \sum_{i=1}^n \ln[1 - G(x_i; \zeta)] + (\alpha - 1) \sum_{i=1}^n \frac{\left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)^\beta \ln \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)}{1 + \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)^\beta} \\ &- \alpha \sum_{i=1}^n \left( 1 + \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)^\beta \right)^{\alpha-1} \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)^\beta \ln \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right), \end{aligned}$$

and

$$\begin{aligned} U_{\zeta_k} &= \sum_{i=1}^n \frac{\partial g(x_i; \zeta) / \partial \zeta_k}{g(x_i; \zeta)} + (\beta - 1) \sum_{i=1}^n \frac{\partial G(x_i; \zeta) / \partial \zeta_k}{G(x_i; \zeta)} + (\beta + 1) \sum_{i=1}^n \frac{\partial G(x_i; \zeta) / \partial \zeta_k}{(1 - G(x_i; \zeta))} \\ &+ \beta(\alpha - 1) \sum_{i=1}^n \frac{\frac{G(x_i; \zeta)^{\beta-1}}{(1 - G(x_i; \zeta))^{\beta+1}} \partial G(x_i; \zeta) / \partial \zeta_k}{1 + \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)^\beta} \\ &- \alpha\beta \sum_{i=1}^n \left( 1 + \left( \frac{G(x_i; \zeta)}{1 - G(x_i; \zeta)} \right)^\beta \right)^{\alpha-1} \frac{G(x_i; \zeta)^{\beta-1}}{(1 - G(x_i; \zeta))^{\beta+1}} \partial G(x_i; \zeta) / \partial \zeta_k. \end{aligned}$$

Setting  $U_\alpha$ ,  $U_\beta$  and  $U_{\zeta_k}$  equal to zero and solving these equations simultaneously yield the maximum likelihood estimate (MLE)  $\hat{\Phi} = (\hat{\alpha}, \hat{\beta}, \hat{\zeta})$  of  $\Phi = (\alpha, \beta, \zeta)^T$ . These equations cannot be solved analytically and statistical software can be used to solve them numerically using iterative methods.

#### 5. Special Models of the OGNH family

In this section, we discuss four special models of the OGNH generated family, namely, OGNH-uniform (OGNHU), OGNH-Lomax (OGNHL), OGNH-Rayleigh (OGNHR) and OGNH-exponential (OGNHE) distributions, respectively. The plots of the pdf and hrf of each special model are also sketched for different parametric values.

##### 5.1. OGNH-Uniform Distribution

Consider the pdf and cdf of the uniformly distributed random variable given by  $g(x; \theta) = \frac{1}{\theta}$ ,  $0 < x < \theta$ , and  $G(x; \theta) = \frac{x}{\theta}$ , respectively. Then the cdf of the OGNHU is given as

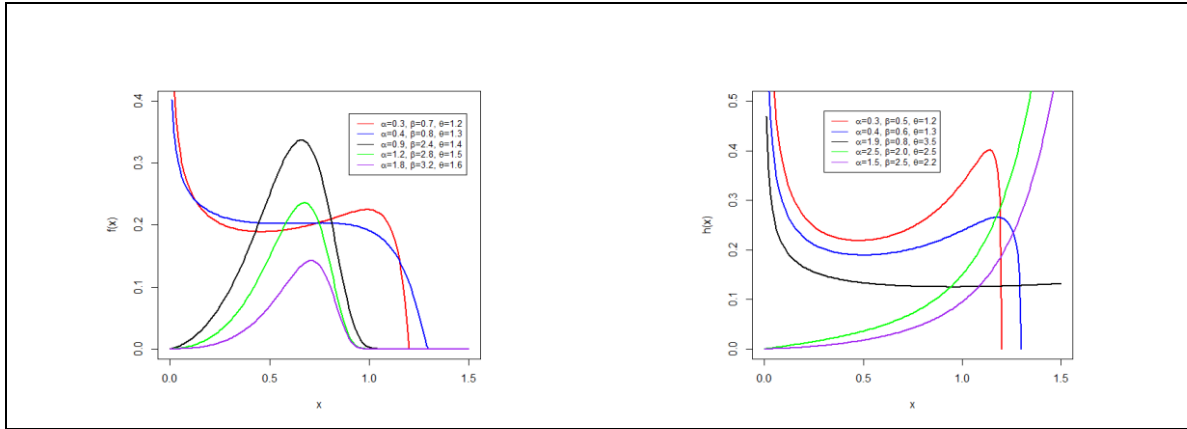
$$F(x) = 1 - \exp \left\{ 1 - \left( 1 + \left( \frac{x}{\theta - x} \right)^\beta \right)^\alpha \right\}, \quad x\alpha, \beta, \theta > 0.$$

The pdf corresponding (3.1) takes the following form



$$f(x) = \frac{\alpha\beta\theta x^{\beta-1}}{(\theta-x)^{\beta+1}} \left(1 + \left(\frac{x}{\theta-x}\right)^\beta\right)^{\alpha-1} \exp\left\{1 - \left(1 + \left(\frac{x}{\theta-x}\right)^\beta\right)^\alpha\right\}.$$

For different values of parameters, the plots of pdf and hrf for the OGNHU distribution are displayed in Figure 1 below:



**Figure 1:** Pdf and hrf of OGNHU distribution for varying values of parameters.

### 5.2. OGNH-Lomax Distribution

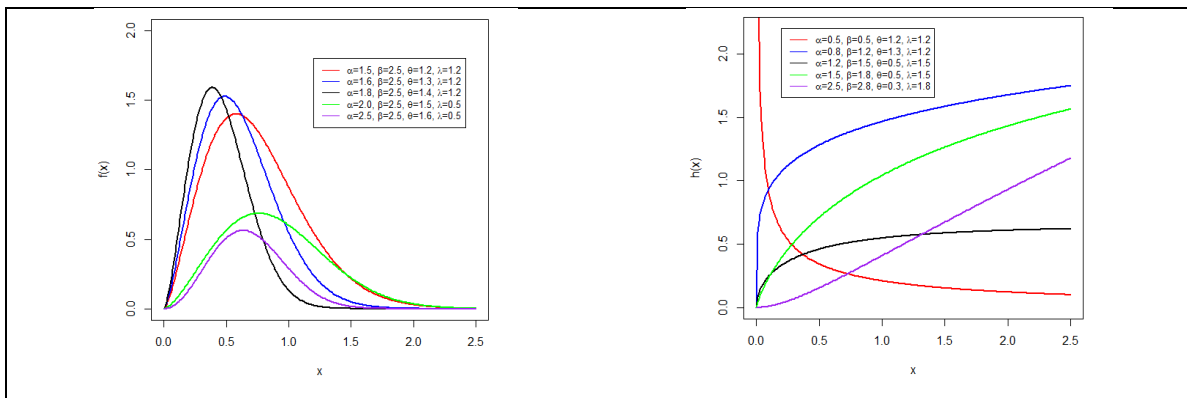
Let the Lomax distribution be the parent distribution with pdf and cdf given by  $g(x) = \theta\lambda(1 + \lambda x)^{-\theta-1}$ ,  $x, \theta, \lambda > 0$ , and  $G(x) = 1 - (1 + \lambda x)^{-\theta}$ , respectively. Then, the cdf and pdf of the OGNHL random variable becomes

$$F(x) = 1 - \exp\left\{1 - \left(1 + ((1 + \lambda x)^\theta - 1)^\beta\right)^\alpha\right\}, \quad x, \alpha, \beta, \theta, \lambda > 0,$$

and

$$f(x) = \frac{\alpha\beta\theta\lambda(1+\lambda x)^{\beta\theta-1}}{(1-(1+\lambda x)^{-\theta})^{1-\beta}} \left(1 + ((1 + \lambda x)^\theta - 1)^\beta\right)^{\alpha-1} \exp\left\{1 - \left(1 + ((1 + \lambda x)^\theta - 1)^\beta\right)^\alpha\right\}.$$

For selected values of the model parameters, the graphs of pdf and hrf for the OGNHL distribution are illustrated in Figure 2 as:



**Figure 2:** Pdf and hrf of OGNHL distribution for selected values of parameters.

### 5.3. OGNH- Rayleigh Distribution

The pdf and cdf of the Rayleigh random variable has the following form  $g(x) = 2\theta x e^{-\theta x^2}$ ,  $x, \theta > 0$ , and  $G(x) = 1 - e^{-\theta x^2}$ , respectively. Then, the cdf of OGNHR distribution is given by

$$F(x) = 1 - \exp\left\{1 - \left(1 + (e^{\theta x^2} - 1)^\beta\right)^\alpha\right\}, \quad x, \alpha, \beta, \theta > 0$$

with pdf

$$f(x) = \frac{2\alpha\beta\theta x e^{\theta\beta x^2}}{(1 - e^{-\theta x^2})^{1-\beta}} \left(1 + (e^{\theta x^2} - 1)^\beta\right)^{\alpha-1} \exp\left\{1 - \left(1 + (e^{\theta x^2} - 1)^\beta\right)^\alpha\right\}.$$

The graphical illustration of the pdf and hrf for the OGNHR are sketched below in Figure 3 as:

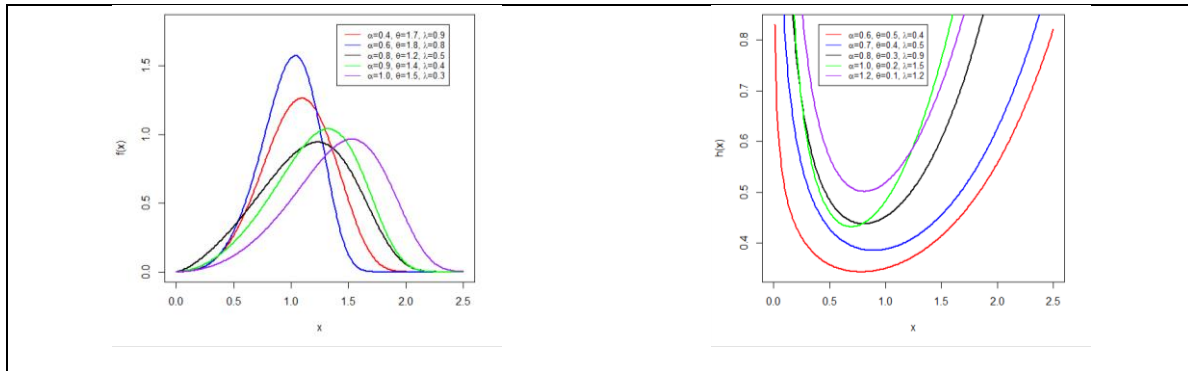


Figure 3: Pdf and hrf of OGNHR distribution for different values of parameters.

### 5.4. OGNH-Exponential Distribution

Considering the exponential distribution as a parent distribution with pdf and cdf given by  $g(x) = \theta e^{-\theta x}$ ,  $x, \theta > 0$ , and  $G(x) = 1 - e^{-\theta x}$ , respectively. Then, the cdf of OGNHE distribution is given by

$$F(x) = 1 - \exp\left\{1 - \left(1 + (e^{\theta x} - 1)^\beta\right)^\alpha\right\}.$$

The density function corresponding to (3.7) becomes

$$f(x) = \frac{\alpha\beta\theta e^{\theta\beta x}}{(1 - e^{-\theta x})^{1-\beta}} \left(1 + (e^{\theta x} - 1)^\beta\right)^{\alpha-1} \exp\left\{1 - \left(1 + (e^{\theta x} - 1)^\beta\right)^\alpha\right\}, \quad x, \alpha, \beta, \theta > 0, \quad (17)$$

The graphical sketch of the pdf and hrf for the OGNHE are showed in Figure 4 as:

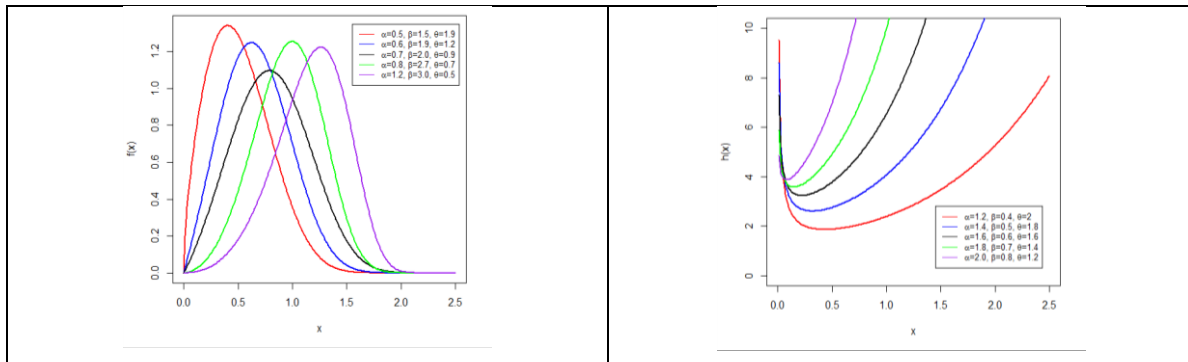


Figure 4: Pdf and hrf of OGNHE distribution for choice values of parameters.

### 5. Properties and Estimation of OGNHE Distribution

In this section, we derived fundamental properties, quantile function, moments, of OGNHE distribution.

#### 5.1. Mathematical and statistical properties

The quantile function of the OGNHE distribution is given by

$$x_u = \frac{1}{\theta} \ln \left[ 1 + \left[ [1 - \ln[1 - u]]^{\frac{1}{\alpha}} - 1 \right]^{\frac{1}{\beta}} \right].$$

Median of the OGNHE random variable is given by

$$M(X) = \frac{1}{\theta} \ln \left[ 1 + \left[ [1 - \ln[0.5]]^{\frac{1}{\alpha}} - 1 \right]^{\frac{1}{\beta}} \right].$$

#### 5.2. Mean deviations

For random variable  $X$  with pdf  $f(x)$ , cdf  $F(x)$ , the mean deviation about the mean and mean deviation about the median, are defined by

$$\delta_1 = 2\mu F(\mu) - 2T(\mu) \text{ and } \delta_2 = \mu - 2T(M),$$

where,  $\mu = E(X)$ ,  $M = \text{Median}(X)$ , and  $T(q) = \sum_{\pi=0}^{\infty} \eta_{\pi} \frac{\gamma(2,q)}{((\pi+1)\theta)^2}$ .

The probability-weighted moments of the OGNHE distribution is given by

$$\tau_{r,s} = \sum_{\pi,v=0}^{\infty} \eta_{\pi} s_v \frac{\Gamma(r+1)}{((\pi+v+1)\theta)^{r+1}}.$$

where

$$\eta_{\pi} = \sum_{k=0}^{\infty} \theta (-1)^{\pi} \binom{\beta(m+1) + k - 1}{\pi} \theta \eta_{\pi, s_v} = \sum_{z=0}^{\infty} (-1)^v \binom{u+z}{v} s_z.$$

The  $r^{\text{th}}$  moment of the OGNHE distribution is given by

$$\mu'_r = \sum_{\pi=0}^{\infty} \eta_{\pi} \frac{\Gamma(r+1)}{((\pi+1)\theta)^{r+1}}.$$

The moment generating function of the OGNHE distribution is given by

$$M_X(t) = \sum_{v,\pi=0}^{\infty} \frac{\eta_{\pi} t^v}{v!} \frac{\Gamma(r+1)}{((\pi+1)\theta)^{r+1}}.$$

The pdf of the  $s^{\text{th}}$  order statistic is given by

$$f_{s:n}(x) = \frac{1}{B(s, n-s+1)} \sum_{v=0}^{n-s} \sum_{\pi,v=0}^{\infty} \eta_{\pi} p_{z,v} e^{-\theta(\pi+v+1)x}.$$

The  $r^{\text{th}}$  moment of  $s^{\text{th}}$  order statistics of OGNHE distribution is given by:

$$E(X_{s:n}^r) = \frac{1}{B(s, n-s+1)} \sum_{v=0}^{n-s} \sum_{\pi,v=0}^{\infty} \eta_{\pi} p_{z,v} \frac{\Gamma(r+1)}{((\pi+v+1)\theta)^{r+1}}.$$

#### 5.3. MLEs and their Performances

The log-likelihood function for parameter vector  $\Phi = (\alpha, \beta, \theta)^T$  is obtained as follows

$$\begin{aligned} \ln L(\Phi) &= n \ln \alpha + n \ln \beta + n \ln \theta + \theta \beta \sum_{i=1}^n x_i + (\beta - 1) \sum_{i=1}^n \ln [1 - e^{-\theta x_i}] \\ &+ (\alpha - 1) \sum_{i=1}^n \ln [1 + (e^{\theta x_i} - 1)^{\beta}] + \sum_{i=1}^n \left\{ 1 - \left( 1 + (e^{\theta x_i} - 1)^{\beta} \right)^{\alpha} \right\}. \end{aligned}$$

The components of score vector are:

$$U_\alpha = \frac{n}{\alpha} + \sum_{i=1}^n \ln [1 + (e^{\theta x_i} - 1)^\beta] - \sum_{i=1}^n (1 + (e^{\theta x_i} - 1)^\beta)^\alpha \ln (1 + (e^{\theta x_i} - 1)^\beta),$$

$$U_\beta = \frac{n}{\beta} + \sum_{i=1}^n \ln [1 - e^{-\theta x_i}] + \theta \sum_{i=1}^n x_i + (\alpha - 1) \sum_{i=1}^n \frac{(e^{\theta x_i} - 1)^\beta \ln(e^{\theta x_i} - 1)}{1 + (e^{\theta x_i} - 1)^\beta} - \alpha \sum_{i=1}^n (1 + (e^{\theta x_i} - 1)^\beta)^{\alpha-1} (e^{\theta x_i} - 1)^\beta \ln(e^{\theta x_i} - 1),$$

and

$$U_\theta = \frac{n}{\theta} + \beta \sum_{i=1}^n x_i + (\beta - 1) \sum_{i=1}^n \frac{x_i e^{-\theta x_i}}{1 - e^{-\theta x_i}} + \beta(\alpha - 1) \sum_{i=1}^n \frac{x_i e^{-\theta x_i} (e^{\theta x_i} - 1)^{\beta-1}}{1 + (e^{\theta x_i} - 1)^\beta} - \alpha\beta \sum_{i=1}^n x_i e^{-\theta x_i} (1 + (e^{\theta x_i} - 1)^\beta)^{\alpha-1} (e^{\theta x_i} - 1)^{\beta-1}.$$

The above equations cannot be solved analytically; rather analytical software is required to solve them numerically. Further, a numerical investigation is carried out to evaluate the performance of ML estimators for *ONHE* model. Performance of estimators is evaluated through their biases, and mean square errors (MSEs) for different sample sizes. A numerical study is performed using Mathematica (9) software. Different sample sizes are considered through the experiments at size  $n = 50, 100, 150$  and  $200$ . In addition, the different values of parameters  $(\alpha, \beta, \theta)$  are considered. The experiment is repeated 10000 times. In each experiment, the estimates of the parameters are obtained by ML method of estimation. The MSEs and biases for the different estimates are reported from these experiments in Table 1.

**Table 1:** The MSEs and Biases of ML estimates for the parameters of the *ONHE* distribution.

<i>n</i>	Par	(0.5,0.5,0.5)			(1.5,0.5,0.5)		
		MLEs	Bias	MSE	MLEs	Bias	MSE
50	$\alpha$	0.5274	0.0274	0.0177	1.5854	0.0854	0.1536
	$\beta$	0.5261	0.0261	0.0167	0.5177	0.0177	0.0096
	$\theta$	0.5199	0.0199	0.0115	0.5080	0.0080	0.0039
100	$\alpha$	0.5142	0.0142	0.0078	1.5456	0.0456	0.0684
	$\beta$	0.5132	0.0132	0.0075	0.5077	0.0077	0.0044
	$\theta$	0.5106	0.0106	0.0056	0.5047	0.0047	0.0020
150	$\alpha$	0.5098	0.0098	0.0051	1.5257	0.0257	0.0451
	$\beta$	0.5091	0.0091	0.0049	0.5042	0.0042	0.0028
	$\theta$	0.5074	0.0074	0.0037	0.5023	0.0023	0.0014
200	$\alpha$	0.5074	0.0074	0.0036	1.5174	0.0174	0.0329
	$\beta$	0.5075	0.0075	0.0035	0.5042	0.0042	0.0021
	$\theta$	0.5057	0.0057	0.0027	0.5014	0.0014	0.0010

**Continued of Table 1**

<i>n</i>	Par	(2,0.5,0.5)			(0.5,1.5,0.5)		
		MLEs	Bias	MSE	MLEs	Bias	MSE
50	$\alpha$	2.1127	0.1127	0.2779	0.5301	0.0301	0.0176
	$\beta$	0.5203	0.0203	0.0120	1.5894	0.0894	0.1548
	$\theta$	0.5064	0.0064	0.0033	0.5492	0.0492	0.0369
100	$\alpha$	2.0533	0.0533	0.1203	0.5136	0.0136	0.0079
	$\beta$	0.5098	0.0098	0.0051	1.5406	0.0406	0.0698
	$\theta$	0.5031	0.0031	0.0016	0.5215	0.0215	0.0137
150	$\alpha$	2.0354	0.0354	0.0813	0.5089	0.0089	0.0047
	$\beta$	0.5052	0.0052	0.0032	1.5264	0.0264	0.0416
	$\theta$	0.5019	0.0019	0.0012	0.5137	0.0137	0.0077
200	$\alpha$	2.0269	0.0269	0.0571	0.5066	0.0066	0.0036
	$\beta$	0.5042	0.0042	0.0023	1.5190	0.0190	0.0303
	$\theta$	0.5016	0.0016	0.0008	0.5102	0.0102	0.0057

**Continued of Table 1**

<i>n</i>	Par	(0.5,2,0.5)			(1.5,1.5,0.5)		
		MLEs	Bias	MSE	MLEs	Bias	MSE
50	$\alpha$	0.5255	0.0255	0.0166	1.5857	0.0857	0.1570
	$\beta$	2.1000	0.0999	0.2580	1.5514	0.0514	0.0849
	$\theta$	0.5737	0.0737	0.1805	0.5770	0.0770	0.1049
100	$\alpha$	0.5132	0.0132	0.0076	1.5447	0.0447	0.0691
	$\beta$	2.0522	0.0522	0.1186	1.5235	0.0234	0.0396
	$\theta$	0.5330	0.0330	0.0248	0.5346	0.0346	0.0229
150	$\alpha$	0.5094	0.0094	0.0048	1.5279	0.0279	0.0438
	$\beta$	2.0363	0.0363	0.0750	1.5163	0.0163	0.0260
	$\theta$	0.5220	0.0220	0.0134	0.5212	0.0212	0.0128
200	$\alpha$	0.5075	0.0075	0.0037	1.5158	0.0158	0.0312
	$\beta$	2.0313	0.0313	0.0570	1.5103	0.0103	0.0188
	$\theta$	0.5168	0.0168	0.0095	0.5129	0.0129	0.0085

**7. Real Life Applications**

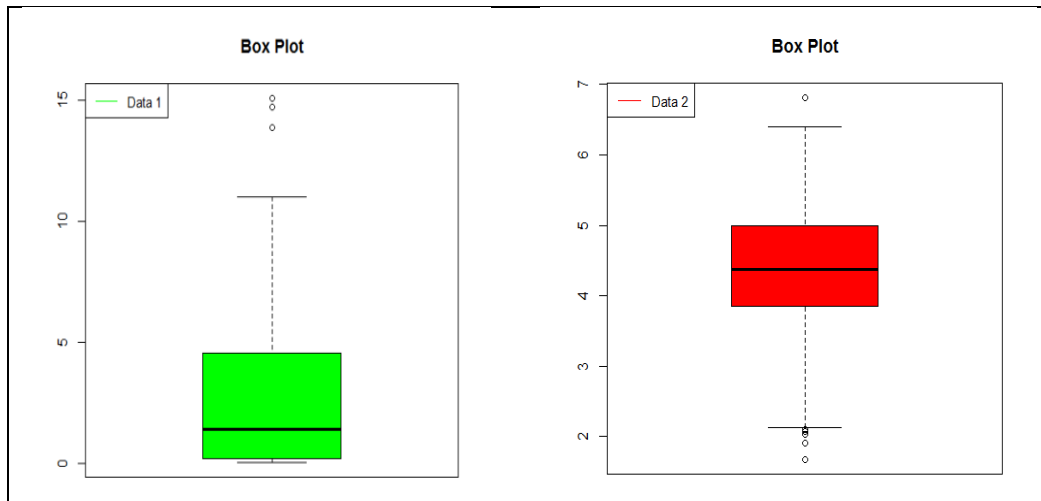
In this section, we have provided two applications using real data sets to illustrate efficiency of the new proposal. The first data set taken from Murthy *et al.* (2004) representing the failure times, while the second data set taken from the website: <http://www.ceramics.nist.gov/srd/summary/ftmain.htm> representing the fracture toughness of Alumina (Al<sub>2</sub>O<sub>3</sub>). For the interest of readers, we have provided the data in Table 2, and summarized in Table 3, as follows

**Table 2:** Failure times and fracture toughness data sets.

Data 1 Failure times data	0.036, 0.058, 0.061, 0.074, 0.078, 0.086, 0.102, 0.103, 0.114, 0.116, 0.148, 0.183, 0.192, 0.254, 0.262, 0.379, 0.381, 0.538, 0.570, 0.574, 0.590, 0.618, 0.645, 0.961, 1.228, 1.600, 2.006, 2.054, 2.804, 3.058, 3.076, 3.147, 3.625, 3.704, 3.931, 4.073, 4.393, 4.534, 4.893, 6.274, 6.816, 7.896, 7.904, 8.022, 9.337, 10.940, 11.020, 13.880, 14.730, 15.080
Data 2 Alumina (Al <sub>2</sub> O <sub>3</sub> )	5.5, 5, 4.9, 6.4, 5.1, 5.2, 5.2, 5, 4.7, 4, 4.5, 4.2, 4.1, 4.56, 5.01, 4.7, 3.13, 3.12, 2.68, 2.77, 2.7, 2.36, 4.38, 5.73, 4.35, 6.81, 1.91, 2.66, 2.61, 1.68, 2.04, 2.08, 2.13, 3.8, 3.73, 3.71, 3.28, 3.9, 4, 3.8, 4.1, 3.9, 4.05, 4, 3.95, 4, 4.5, 4.5, 4.2, 4.55, 4.65, 4.1, 4.25, 4.3, 4.5, 4.7, 5.15, 4.3, 4.5, 4.9, 5, 5.35, 5.15, 5.25, 5.8, 5.85, 5.9, 5.75, 6.25, 6.05, 5.9, 3.6, 4.1, 4.5, 5.3, 4.85, 5.3, 5.45, 5.1, 5.3, 5.2, 5.3, 5.25, 4.75, 4.5, 4.2, 4, 4.15, 4.25, 4.3, 3.75, 3.95, 3.51, 4.13, 5.45, 2.1, 4.6, 3.2, 2.5, 4.1, 3.5, 3.2, 3.3, 4.6, 4.3, 4.3, 4.5, 5.5, 4.6, 4.9, 4.3, 3, 3.4, 3.7, 4.4, 4.9, 4.9, 5

**Table 3:** Descriptive measures of the data given in table 1.

Summary Statistics	Min.	1 <sup>st</sup> Quartile	Median	Mean	3 <sup>rd</sup> Quartile	Max.
Data 1	0.0360	0.2075	1.4140	3.3430	4.4990	15.0800
Data 2	1.680	3.850	4.380	4.325	5.000	6.810



**Figure 5:** Box plots of the data given in table 1.

We compare the efficiency of the new proposal to some well-known lifetime distributions, namely; generalized linear exponential (GLE), exponentiated generalized linear exponential (Ex-GLE), exponential Lomax (EL), alpha power exponential (APE) and alpha power transformed Weibull (APTW) distributions. The cdfs of the other fitted models are

Generalized linear exponential of Mahmoud and Alam (2010)

$$G(x) = 1 - e^{-(\alpha x^2 + \beta x)^\theta} \quad x, \alpha, \beta, \theta > 0.$$

Exponentiated generalized linear exponential by Sarhanet al. (2013)

$$G(x) = \left(1 - e^{-(\alpha x^2 + \beta x)^\theta}\right)^\lambda \quad x, \alpha, \beta, \theta, \lambda > 0.$$

Exponential Lomax studied proposed by El-Bassiouny et al. (2015)  $G(x) = 1 - e^{-\theta\left(\frac{\beta}{x+\beta}\right)^{-\alpha}}$ ,  $x \geq -\beta, \alpha, \beta, \theta > 0$ .

Alpha power exponential studied by Mahdavi and Kundu (2017)

$$G(x) = \frac{\alpha^{(1-e^{-\beta x})} - 1}{\alpha - 1}, \quad x, \alpha, \beta > 0, \alpha \neq 1.$$

Alpha power transformed Weibull introduced by Dey et al. (2017)

$$G(x) = \frac{\alpha^{(1-e^{-\beta x^\theta})} - 1}{\alpha - 1}, \quad x, \alpha, \beta, \theta > 0, \alpha \neq 1.$$

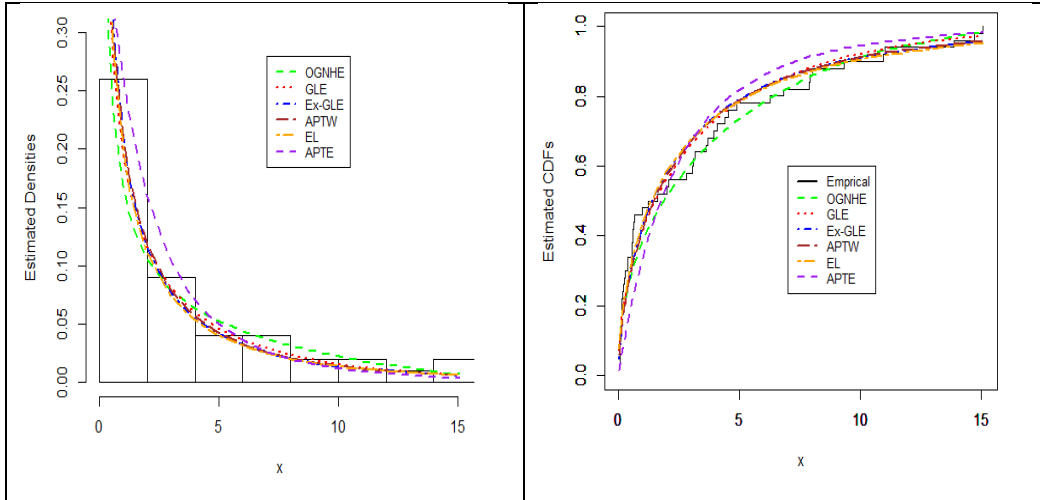
To show the goodness of fit of the fitted models, we consider the analytical measures including Kolmogorov–Smirnov (KS) statistic, Anderson Darling (AD) statistic, Cramer-Von-Misses (CM) statistic, log-likelihood (LL), Akaike Information Criterion (AIC), Bayesian information criterion (BIC), corrected Akaike information criterion (CAIC) and Hannan-Quinn information criterion (HQIC). The analytical measure are given in Table 4, 5, 6 and 7 shows that the suggested method provides best fit than the other fitted distributions.

**Table 4:** Maximum likelihood estimates and their respective standard errors in the parenthesis for data 1.

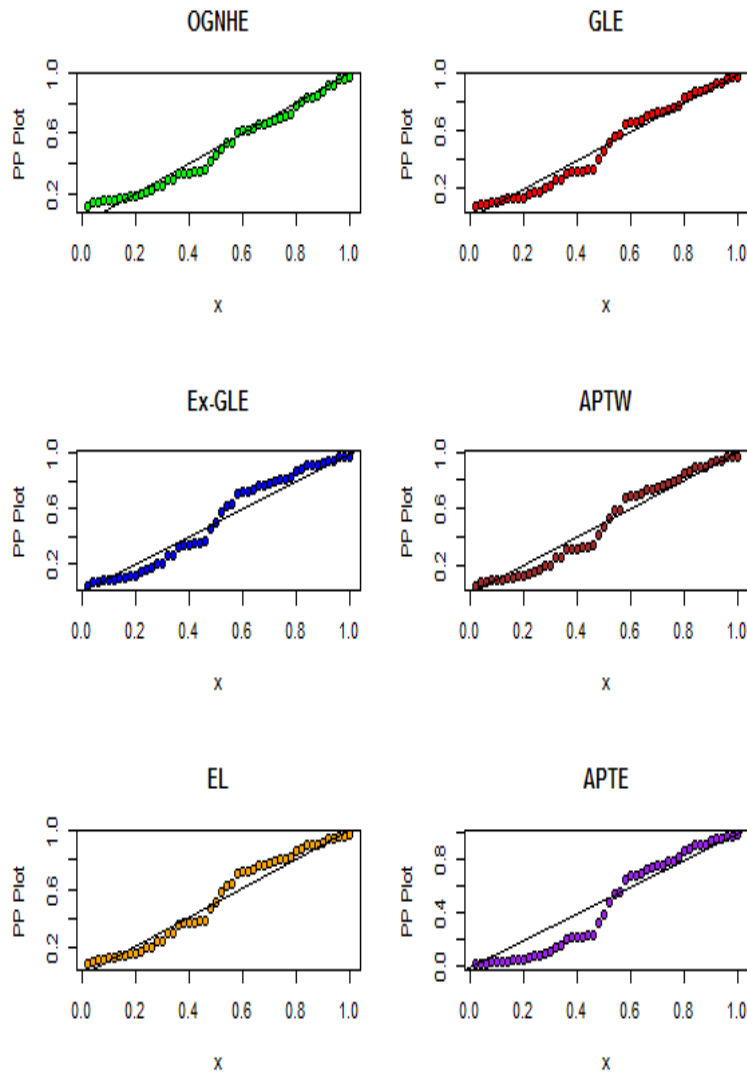
Dist.	$\hat{\alpha}$	$\hat{\beta}$	$\hat{\theta}$	$\hat{\lambda}$
<b>OGNHE</b>	0.481 (0.4013)	0.497 (0.0632)	0.148 (0.0859)	
<b>GLE</b>	0.015 (0.018)	0.535 (0.1170)	0.589 (0.0915)	
<b>APTW</b>	0.798 (1.164)	0.500 (0.2626)	0.669 (0.1146)	
<b>Ex-GLE</b>	6.578 (10.829)	0.688 (1.462)	0.274 (0.0758)	4.287 (2.5937)
<b>EL</b>	0.609 (0.0647)	0.0037 (0.00146)	0.190 (0.0095)	
<b>APTE</b>	0.531 (0.1336)	0.589 (0.0629)		

**Table 5:** Analytical results of the proposed and other competing models for data 1.

Dist.	KS	AD	CM	LL	AIC	BIC	CIAC	HQIC
OGNHE	0.139	0.871	0.129	101.52	209.05	214.79	209.57	211.24
GLE	0.122	0.885	0.131	101.77	209.54	215.28	210.07	211.73
APTW	0.140	0.892	0.133	102.12	210.64	216.38	211.16	212.83
Ex-GLE	0.115	0.896	0.139	102.28	210.57	217.22	211.46	213.48
EL	0.142	0.975	0.155	102.96	211.93	217.66	212.45	214.11
APTE	0.220	1.027	0.165	107.09	218.19	222.01	218.45	219.65



**Figure 6:** Estimated pdfs and cdfs of the fitted models corresponding to data 1.



**Figure 7:** PP plots of the fitted models corresponding to data 1.

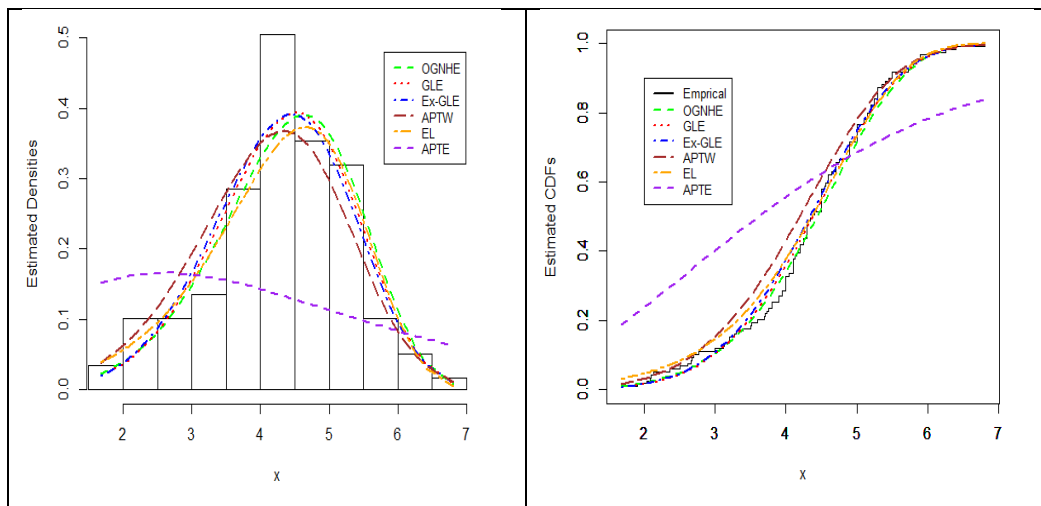


**Table 6:** Maximum likelihood estimates and their respective standard errors in the parenthesis for data 2.

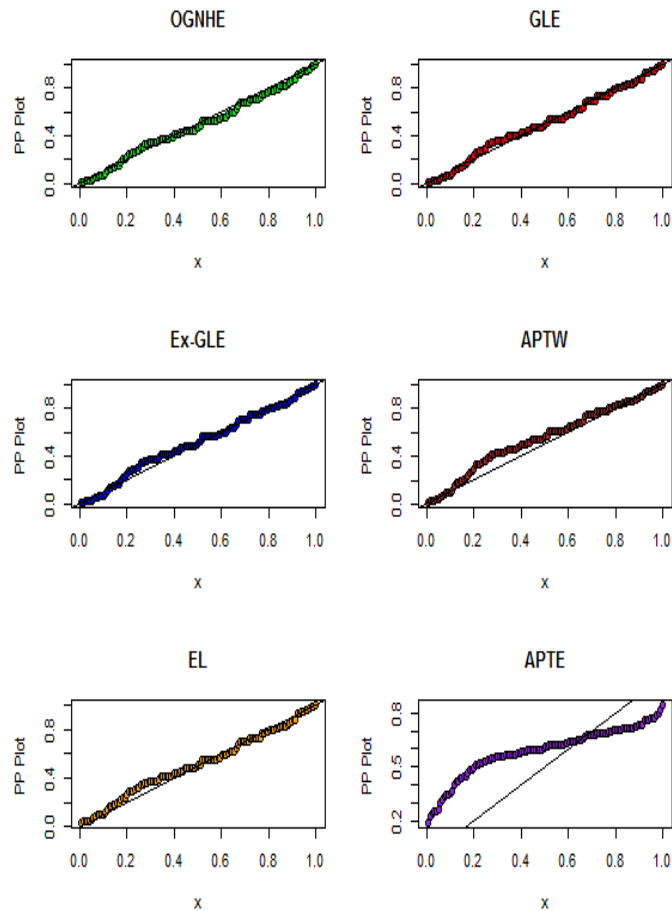
Dist.	$\hat{\alpha}$	$\hat{\beta}$	$\hat{\theta}$	$\hat{\lambda}$
OGNHE	0.974 (0.3625)	0.641 (0.4203)	0.1462 (0.0159)	
GLE	0.0420 (0.5398)	0.0542 (0.0118)	2.0655 (0.3070)	
APTW	2.3290 (0.7308)	0.0329 (0.0789)	3.9312 (0.1318)	
Ex-GLE	0.0064 (0.1245)	3.0452 (0.5231)	2.3923 (0.5983)	1.0535 (0.4834)
EL	14.7402 (3.1329)	9.8520 (2.2892)	0.0039 (0.0059)	
APTE	21.442 (6.488)	0.4295 (3.0274)		

**Table 7:** Analytical results of the proposed and other competing models for data 2.

Dist.	KS	AD	CM	LL	AIC	BIC	CIAC	HQIC
OGNHE	0.063	0.473	0.078	168.64	343.28	351.62	343.49	346.67
GLE	0.092	0.516	0.802	168.70	345.17	353.51	345.38	348.56
APTW	0.144	0.573	0.087	170.60	347.21	355.55	347.42	350.60
Ex-GLE	0.087	0.565	0.084	168.84	345.68	354.79	346.03	350.19
EL	0.093	0.591	0.149	170.82	347.65	355.99	347.86	351.03
APTE	0.310	2.041	0.334	245.59	495.18	500.74	495.28	497.44



**Figure 8:** Estimated pdfs and cdfs of the fitted models resultant to data 2.



**Figure 9:** PP plots of the fitted models corresponding to data 2.

### 8. Conclusion

In this paper, we introduce the odd generalized NH-G family. Some of its properties are derived and some members of the family are defined. A member of the odd generalized NH -G family, namely, the odd generalized NH exponential distribution is defined and studied. Various properties of the odd generalized NH exponential distribution including, probability weighted moments, moments, Rényi entropy and order statistics are derived. Estimation and simulation issues for the *OGNHE* model are performed to study the behavior of the estimated parameters. The *OGNHE* is applied to fit a real data set. This application shows that the *OGNHE* can be preferred over some other well-known distributions.

### Appendix A

Theorem 1. Let  $(\Omega, F, P)$  be a given probability space and let  $H = [a, b]$  be an interval for some  $d < b (a = -\infty, b = \infty$  might as well be allowed).

Let  $X: \Omega \rightarrow H$  be a continuous random variable with the distribution function  $F$  and let  $q_1$  and  $q_2$  be two real functions defined on  $H$  such that

$$E[q_2(x)/X \geq x] = E[q_1(x)/X \geq x]\xi(x) \quad X \in H,$$

is defined with some real function  $\eta$ . Assume that  $q_1, q_2, C^1(H), \xi \in C^2(H)$  and  $F$  is twice continuously differentiable and strictly monotone function on the set  $H$ . Finally, assume that the equation  $\xi q_1 = q_2$  has no real solution in the interior of  $H$ . Then  $F$  is uniquely determined by the functions  $q_1, q_2$  and  $\xi$ , particularly.

$$F(x) = \int_a^x C \left| \frac{\xi'(u)}{\xi(u)q_1(u) - q_2(u)} \right| \exp(-S(u)) du,$$

Where the function  $S$  is a solution of the differential equation  $S' = \frac{\xi'/q_1}{\xi q_1 - q_2}$  and  $C$  is the normalization constant, such that  $\int_H dF = 1$ .

We like to mention that this kind of characterization based on the ratio of truncated moments is stable in the sense of weak convergence (see, Glänzel [2]), in particular, let us assume that there is a sequence  $\{X_n\}$  of random variables with distribution functions  $\{F_n\}$  such that the functions  $q_{1n}, q_{2n}$  and  $\xi_n, (n \in \mathbb{N})$  satisfy the conditions of Theorem 1 and let  $q_{1n} \rightarrow q_1, q_{2n} \rightarrow q_2$  for some continuously differentiable real functions  $q_1$  and  $q_2$ . Let, finally,  $X$  be a random variable with distribution  $F$ . Under the condition that  $q_{1n}X_n$  and  $q_{2n}X_n$  are uniformly integrable and the family  $\{F_n\}$  is relatively compact, the sequence  $X_n$  converges to  $X$  in distribution if and only if  $\xi_n$  converges to  $\xi$ , where

$$\xi(x) = \frac{E[q_2(x)/X \geq x]}{E[q_1(x)/X \geq x]}.$$

This stability theorem makes sure that the convergence of distribution functions is reflected by corresponding convergence of the functions  $q_1, q_2$  and  $\xi$ , respectively. It guarantees, for instance, the 'convergence' of characterization of the Wald distribution to that of the Lévy-Smirnov distribution if  $\alpha \rightarrow \infty$ .

A further consequence of the stability property of Theorem 1 is the application of this theorem to special tasks in statistical practice such as the estimation of the parameters of discrete distributions. For such purpose, the functions  $q_1, q_2$  and, specially,  $\xi$  should be as simple as possible. Since the function triplet is not uniquely determined it is often possible to choose  $\xi$  as a linear function. Therefore, it is worth analyzing some special cases which helps to find new characterizations reflecting the relationship between individual continuous univariate distributions and appropriate in other areas of statistics.

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