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# The Political Logic of Government Disclosure: Evidence from Information Requests in Mexico

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When citizens ask questions, how does their government answer? Requests for government information confront officials with incentives both for and against disclosure. We argue that officials seek to manage political risks in ways that favor requests from government-aligned regions. We study responsiveness in the context of Mexico's access-to-information law, using publicly available data from several hundred thousand information requests filed with Mexican federal government agencies between 2003 and 2015. Our empirical strategy makes comparisons only among requests sent to similar agencies on similar topics at similar times, while accounting for the complexity, sophistication, and sensitivity of individual requests. We find that requests filed from locales with higher governing-party vote shares receive more favorable responses, across multiple indicators of the nature and timing of responses. Further, we find bias only for requests on publicly relevant topics, providing evidence in favor of a mechanism of mitigating political risks over one of rewarding supporters with greater access to benefits.

To whom are agents of accountability themselves accountable? Government officials charged with responding to citizen queries and disclosing official information play a fundamental role in downward "vertical accountability," in which the state answers to citizens for its actions (O'Donnell 1998; Schedler 1999). And yet, as agents of political principals—both their hierarchical superiors and elected politicians—they may face conflicting incentives to also be accountable upward, in ways that inhibit the provision of information to citizens. How do government officials resolve these roles, and with what effects on responsiveness to citizens?

While many scholars have studied how electorally motivated discretion drives government expenditures or public goods provision (Besley and Burgess 2002; Carlitz 2017; Diaz-Cayeros, Estevez, and Magaloni 2016; Golden and Min 2013;

Kramon and Posner 2013; Stokes et al. 2013), government responsiveness extends beyond spending decisions: governments interface with citizens in millions of everyday interactions, in which agents of the state may vary in their responsiveness to citizen needs. We examine government responsiveness in one case of frequent citizen-government interaction: responses to citizen requests for government information in Mexico. To do so, we use a novel data set incorporating every information request filed with Mexican federal government agencies between 2003 and 2015.

Much like government expenditures, government information can be divided into particularistic and public categories. For instance, information useful in applying for government benefits or employment serves the requester's private interest, while information that reveals government corruption

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or inefficiency serves the public's interest in good governance. While this categorization mirrors the public-private goods distinction made in the distributive politics literature, the provision of government information produces a different set of incentives for politicians than does distributive spending, in particular because publicly relevant information can present a political risk to the incumbent administration.

We thus focus attention on the incentives at work in the specific domain of government information. Officials charged with responding to citizens serve two distinct roles that are central to democratic governance and accountability. First, they are gatekeepers for information about government activities—from budgets to security to evidence of corruption—relevant for citizens' retrospective evaluations of government performance, a central component of democratic accountability (Przeworski, Stokes, and Manin 1999). Second, they provide information useful for guiding citizens' everyday dealings with government agencies, such as applying for benefits, acquiring titles or permits, or seeking government employment. As the timely access to such information confers material advantages, it can become a currency of clientelism if granted discretionarily according to partisanship (Stokes et al. 2013, 7). Following from these two roles, discretionary nondisclosure of public information can either challenge political accountability or can hinder the quality of governance by violating the principle of impartiality (Rothstein and Teorell 2008) in the provision of public goods and services. It is the goal of this paper to assess these two forms of discretionary disclosure.

For the agents of accountability at the heart of this process, motivations for responding positively include commitment to the goals of transparency as a tool for downward accountability, as well as avoiding formal or informal sanctions through state institutions of horizontal accountability. Such agents may also be motivated by partisan incentives to provide information that helps political allies access state benefits or navigate regulatory processes. However, they also have many reasons to be less responsive, particularly concerning the risk of unwanted disclosures resulting in scandal or negative media attention that might harm the reputation of their agency's leadership or the political goals of the ruling party—both forms of upward accountability.

How do officials manage these competing incentives regarding disclosure of government information? We argue that officials filter and discriminate based on the perceived political risk of requests, informed by the partisan alignment of communities where requests originate. However, counter to the distributive politics literature, which expects discretionary targeting of private goods, we argue that officials prioritize areas of core support in the provision of information that is most publicly relevant. As the latter type of information presents a

greater risk of generating negative attention, politicians and their agents prefer to withhold this information from political opponents where possible. We thus suggest that government responsiveness in the domain of information, rather than spending or material goods, is better understood through such a logic of managing risks rather than a logic of targeting benefits.

We test these expectations by comparing the quality of responses across requests originating from municipalities with varying levels of support for the governing party—but filed through the same centralized and ostensibly anonymous system. We address heterogeneity in requests originating in areas of high and low governing-party support by considering variation only among requests sent to the same government agency, on the same topic, and in the same year, and by incorporating text-based information on the complexity, sensitivity, legalism, and sophistication of individual requests. To control for the topics underlying requests, we use unsupervised topic models to classify requests into twenty topics, which we categorize as either primarily publicly relevant (e.g., budgets and spending, security issues, and environmental disputes) or primarily related to private goals (e.g., accessing services and benefits, or seeking jobs or contracts).

We find evidence that requests filed from areas of stronger government support receive more favorable responses: faster, less likely to claim the information does not exist, and less likely to raise financial or physical barriers. These results are stronger among publicly relevant topics than among topics with a more private focus, consistent with a mechanism of officials managing political risk, and inconsistent with a mechanism of rewarding allies with preferential access to material benefits. Further analyses reveal evidence of within-municipality variation in responsiveness over time, exploiting changes in political alignment after the Institutional Revolutionary Party (PRI) regained presidential power in 2012.

Using qualitative evidence, we surmise that bias in responsiveness to information requests is primarily due to officials inferring the political risk posed by a given request based on the traits of the municipality from which it originates—information that is often readily available to the bureaucrat. However, in some cases officials may be able to more precisely infer the identity of the requester and thus discriminate against political opponents or notable watchdog groups. Observationally, this more fine-grained discretion would yield the same results that we observe: worse responses to information requests that come from areas where political opponents of the ruling party predominate.

Interviews with bureaucrats at several liaison units charged with handling information requests in government agencies illustrate that the types of discretion we elucidate are procedurally

feasible and clarify the conflicting incentives presented by requests for government information. While Mexico's access-to-information (ATI) system was designed to "level the playing field" by anonymizing requests, liaison unit personnel report several informal avenues by which requester traits can potentially become known and influence responsiveness. Further, officials described both incentives favoring disclosure—such as avoiding sanctions for noncompliance or commitment to the goals of the transparency law—as well as considerations of the political risk that disclosure may pose to their supervisors and political principals, particularly for sensitive topics.

While we expect similar political dynamics to be at work in countries other than Mexico, this context offers several advantages for our analysis. On a practical level, the online request system offers richly detailed data over a long period of time. More crucially, as a federal democracy with stable multiparty competition, Mexico presents significant regional variation in political dynamics. Moreover, while Mexico has been at the vanguard among transitional democracies in the design and implementation of ATI institutions, persistent corruption and bureaucratic discretion introduce both motive and opportunity for biased responsiveness to information requests. However, we expect that our findings would extend to other contexts where bureaucrats have discretion in disclosing government information, yet face political risks in doing so.

### GOVERNMENT RESPONSIVENESS AND INFORMATION DISCLOSURE

We combine insights from two literatures on government responsiveness. First, studies of policy responsiveness address the extent to which government policies respond to voter preferences (Golden and Min 2013). Such studies focus on partisan incentives that shape government spending and public goods provision, typically across subnational units (Besley and Burgess 2002; Cleary 2007; Magaloni 2006). A central debate in this literature regards how politicians allocate benefits between areas of ruling- and opposition-party strength (Diaz-Cayeros et al. 2016; Gans-Morse, Mazzuca, and Nichter 2014; Stokes et al. 2013).

A second group analyzes individual responsiveness, the extent to which individual citizen needs are met in everyday interactions with the state. Recent studies have identified bias regarding citizens' inferred race, ethnicity, gender, and social status (Butler and Broockman 2011; Distelhorst and Hou 2014; McClendon 2016; White, Nathan, and Faller 2015). Studies of responsiveness to information requests, the present focus, have identified differences based on the threat of collective action in China (Chen, Pan, and Xu 2016) and the design of government agencies in the United States (Wood and Lewis 2017). And studies of ATI in Mexico specifically have identi-

fied limited favoritism of citizens with implied political connections (Lagunes and Pocasangre 2019)<sup>1</sup> and biases against citizens seeking information about the drug war (Almanzar, Aspinwall, and Crow 2018). Yet we lack understanding of how broader partisan objectives shape bureaucratic responsiveness in interactions with individual citizens.

ATI institutions enable citizens to request information from public officials, and obligate officials to respond, subject to certain legal exemptions. In practice, officials operate in conditions of limited capacity (in terms of staff, finances, and records-management systems), substantial discretion, and inconsistent oversight, making ATI institutions a promising venue to study responsiveness. Further, they generate data about responsiveness across two quite different areas of citizen activity: private goals for the benefit of the requester and public goals related to governance and accountability (Michener and Worthy 2018). Information requests in the former category include, for example, questions about how to access government benefits or programs, conduct bureaucratic procedures, or seek government contracts or jobs. Public goals, on the other hand, include investigating corruption, patronage, or other wrongdoing, evaluating decision-making or policy effectiveness, and mobilizing for political action. Similar to studies that find different targeting logics for different types of distributive goods (Albertus 2012; Diaz-Cayeros et al. 2016; Kramon and Posner 2013; Palmer-Rubin 2016), we postulate that public and private varieties of information generate different incentive structures for officials.

As shown in table 1, certain incentives favor responsiveness across both public and private information. Disclosure of requested information provides citizens a positive impression of government performance. Officials may feel committed to comply with legal mandates or the ethos of transparency, or may wish to avoid sanction for nondisclosure, either within their agencies or from oversight bodies. Officials also have reasons to be unresponsive to all types of information requests. Aside from standard criteria for withholding information (confidentiality, national security, commercial secrets, etc.), bureaucrats may be discretionarily unresponsive out of an interest in minimizing workloads, particularly in understaffed agencies.

However, other potential incentives, denoted in italics, are specific to the type of information. A core consideration in disclosing public information is political risk: the potential of the requested information to damage the reputation of the agency, its leadership, or the governing party. Political elites from the ruling party clearly wish to avoid scandals, investigations, or information revelations in the media that might

1. Michener et al. (2020) find similar results in Brazil.

Table 1. Incentives For and Against Information Disclosure

	Incentives for Disclosure	Incentives against Disclosure
Information for public goals	Demonstrating performance; Personal commitment; Avoiding sanctions	Limiting workload; Complying with legal exemptions; <i>Managing political risks</i>
Information for private goals	Demonstrating performance; Personal commitment; Avoiding sanctions; <i>Targeting benefits</i>	Limiting workload; Complying with legal exemptions

Note. Potential incentives that are specific to the type of information requested are denoted in italics.

harm themselves or their allies. Agency leaders also seek to protect their own future political careers. These considerations weigh against disclosure for topics of information that pose greater political risks. In contrast, for information related to private goals, disclosure might follow a logic of targeting benefits, as does distributive spending, strategically disclosing information that helps the requester get a job, win a contract, or receive government benefits, in order to reward supporters and punish opposition.

Officials responding to information requests consider both their own interests (including individual career goals and organizational reputation) and the interests of their political principals. The prevalence of political appointees and patronage in bureaucratic agencies produces dependence on the governing political party's interests. We thus expect that officials responding to information requests operate, at least in part, as agents of the party controlling the presidency, and act in accordance with their political interest.

Both public and private goals can be furthered by favoring copartisans. Officials can discriminate directly in cases of known requesters, or use the political alignment of the location from which a request originates to infer the likelihood that the requester is an "ally" or "adversary." Either case means that officials respond in ways that disproportionately favor government supporters and government-supporting areas. Thus, our Partisan Discretion hypothesis predicts that, across all types of requests, those from government-supporting areas receive more favorable responses than those from opposition-supporting areas.

We further pose two subsidiary and competing hypotheses, regarding whether such partisan discretion follows a logic of targeting benefits or one of mitigating political risk. A naive deployment of findings from the distributive politics literature would predict a strategy based on targeting benefits. As such, if officials manipulate disclosure decisions as part of a clientelistic strategy to target voters with particularistic benefits (Diaz-Cayeros et al. 2016; Gans-Morse et al. 2014; Stokes et al. 2013), we would expect to see stronger effects for topics of information requests that contribute to private, rather than

public goals. The Targeting Benefits hypothesis predicts that partisan effects are stronger for requests on topics of information about private needs or benefits than about information with public potential.

In light of the above discussion about the acute political risk posed by disclosure of information with public relevance, however, we have reason to expect the opposite. Given the high costs associated with scandal weighed against the marginal payoff for helping allies access benefits—as well as incentives to appear impartial and maintain high response rates—we hypothesize that officials may exercise more discretion in disclosing information that poses political risks than in disclosing that which generates private benefits. Therefore, the Mitigating Risk hypothesis predicts that partisan effects are stronger for requests on topics of information with public potential than about private needs or benefits.

#### ACCESS TO INFORMATION IN MEXICO

Mexico's Ley Federal de Transparencia y Acceso a la Información Pública Gubernamental (LFTAIPG) was passed in June 2002, taking effect one year later. An ATI law was a campaign promise of President Vicente Fox's Partido Acción Nacional (PAN) in the watershed 2000 presidential election. Although Fox initially proposed weak draft legislation, sustained civil society, media, and opposition pressure contributed to the ultimate passage of a law with cross-party support, widely hailed as one of the world's strongest (Berliner and Erlich 2015; Michener 2011; Pinto 2009). The law's implementation and wide usage is seen as a model among developing countries (Bookman and Guerrero Amparán 2009; Michener 2011).

Notably, the ATI law created an independent information commission, the Instituto Federal de Acceso a la Información (IFAI),<sup>2</sup> tasked with promoting the new law, monitoring compliance, and resolving appeals. IFAI also created a centralized

2. In 2015, its name changed to INAI.

online information portal through which government agencies manage requests. Citizens file requests and receive responses almost exclusively (97% of all requests) through this portal, known as INFOMEX until 2016. When citizens do file handwritten requests (only 3% of all cases), agency officials manage these and enter the relevant information into the portal's database.

When a request is filed, it is first sent from the INFOMEX system to the *unidad de enlace*, or transparency liaison unit, within the government agency to which the request was directed. The unit's staff reviews the request and directs it to the relevant administrative units within the agency, unless simple or common enough that it can be handled directly. Administrative units then return preliminary responses to the liaison unit, with further communication if necessary before responses are returned via INFOMEX. Most agencies manage these internal processes using combinations of e-mails and spreadsheets, with only 10% reporting automated management systems even as late as 2016 (Ríos Cázares, Castañeda, and García 2017). Each agency also has an internal information committee that reviews and approves some categories of denials or deadline extensions. The head of the liaison unit sits on the information committee, along with a representative of the agency head and an internal control officer.<sup>3</sup> This creates a formal venue for discussions around how to respond, although informal discussions can also take place.

While several observers claim that INFOMEX maintains requester anonymity (Guerrero Amparán and Toledo 2009; Mizrahi and Mendiburu 2012; Sobel et al. 2006) this is not generally the case. Indeed, Lagunes and Pocasangre (2019, 167) observe that "while Mexico's FOIA system is supposed to be identity blind, we rarely see instances in which the identity of the individual is blacked out from requests." And Fox and Haight (2007, 40) note that "at first, the IFAI tried to prevent agencies from accessing personal data in SISI [INFOMEX's earlier name] accounts, however agencies insisted on knowing who was requesting information. The IFAI accepted this position, in the absence of a legal basis for defending the anonymity of requestors."

Individuals must create an account to use INFOMEX, and provide not only their name and contact information but also their physical address. This requester information can be seen by liaison unit staff who first receive requests and distribute them within their agency. Some agencies use specialized software to keep requester information out of internal communications, but most do not. A 2016 survey of liaison units found

that 99% reported communicating about requests over e-mail and 86% over telephone (Ríos Cázares et al. 2017, 61). Even where specialized software exists, however, anonymity is not guaranteed. In one large, cabinet-level ministry, liaison unit staff claimed that their internal system did allow other officials to see full requester information.<sup>4</sup>

This complicated situation regarding requester information leaves several avenues for discrimination, either by specific requester identity (which would only sometimes be informative, and which we cannot observe in the publicly available data) or by location (which is both systematically informative and which we can observe). First, officials in the liaison unit—who can formally see requester information—might alter the speed or nature of initial decisions based on the requester. Second, the agency's information committee might make different decisions regarding extensions or claims of classification, confidentiality, or nonexistence of the requested information. Third, other officials may see requester information despite mandates to the contrary, due to the agency's internal procedures or informal pressure. Several interviewees suggested that other officials indeed attempt to find out information about the requester (and potentially incorporate such information into their decisions) and are sometimes successful.

There is a clear incentive for selective disclosure, as information requests often lead to public scandals or negative media coverage. For instance, newspaper stories frequently reference the use of information requests in reporting and investigation. These include information on potential misuse of funds (*Reforma* 2005, 2015a) public employees lacking qualifications (*Reforma* 2014), irregularities in procurement or handling of supplies (*Reforma* 2008; *El Universal* 2013), and even to investigate potential arbitrary detention by police forces and the discovery of mass graves (*Reforma* 2012, 2015b).

In managing the political risks of disclosure, officials face conflicting incentives. Interviewed liaison unit officials emphasized both their commitment to the law and to serving citizens, and their fear of "getting in trouble" with IFAI, or of attracting negative attention to the agency for poor compliance records—all incentives in favor of disclosure. But officials also discussed political considerations. One official said that, for some sensitive requests (often relating to patronage spending), higher-level officials simply order the denial of the request, leaving the liaison unit to come up with a defensible justification. Another said that agency leadership sometimes argue directly with the information committee in favor of denying specific requests. The head of yet another liaison unit directly invoked

3. Several procedural details were changed by a 2016 reform, but this article refers to the situation during the period under study.

4. Interviews were conducted with officials in the liaison units of seven different government agencies in March 2017, under Arizona State University IRB number 00005773.

the reputation of the head of the agency, saying “we have to take care of the secretary. . . . We have to protect him.”

A 2007 survey (Díaz Iturbe 2007) of 1,221 high level staff at 150 federal entities further demonstrates these mixed incentives. The survey found 73% agreeing that “the federal government should give access to all the information that it generates” (37) and 76% aware of sanctions for noncompliance, 29% of whom cited suspension as a likely penalty (26). Yet at the same time, 53% reported that they were justified in withholding requested information when they suspected the information could be used to attack the agency (43), and respondents were broadly aware of a variety of strategies to withhold information, such as declaring it nonexistent or the responsibility of another agency, or requiring the requester to collect a physical copy (44).

While government agencies and their officials do not directly face electoral incentives, they cannot remain fully independent of them either. Top agency leadership are appointees accountable directly to the President, and often seek to protect their own reputations in pursuit of future political advancement. Liaison unit heads are frequently replaced, although in some cases they remain in place through presidential transitions (even in 2012). Lower-level staff within each unit, particularly those with substantial experience or legal expertise, are more likely to remain in place through a transition. Nearly all relevant staff to this process are *empleados de confianza*—technically patronage appointees who may be fired at will, despite usually pursuing longer-term careers in the public sector, and thus face incentives to protect their principals’ reputations. One interviewed official even noted liaison unit directors had been fired in the past for disclosing information that displeased agency leadership.

A 2016 survey of liaison units (Ríos Cázares et al. 2017, 58–61) found that 88% of unit directors (and 76% of responsible staff) were *empleados de confianza* (with the bulk of the remainder being even less secure temporary employees), averaging 91 (93) months in the agency, and 33 (57) months in the response unit, respectively. Nearly all had at least undergraduate degrees, while roughly 40% of directors had either master’s or doctorate degrees. The average response unit had four to five staff, including the director. A survey of officials at the ATI system’s founding (Gill and Hughes 2005, 132) found that “nearly one-half of respondents thought government officials would only partially comply with the new access law and said potential sanction by superiors, or the fear of such sanction, would prompt non-compliance.” Furthermore, given a policy requiring that unit directors report directly to the agency head, these ATI authorities are usually high-level agency officials, invested in the well-being of the agency and its leadership (Ríos Cázares et al. 2017, 52).

## SAMPLE, PREPROCESSING, AND TEXT ANALYSIS

We focus on Mexico’s publicly available administrative data containing all ATI requests filed with federal government agencies during the period from June 2003 (when the law came into effect) through August 2015, as made available by the Mexican government via INFOMEX.<sup>5</sup> These records contain detailed information on each request, including the date and time filed, the agency to which the request is directed, the full request text, the date of response, the nature of the response, and links to attached files associated with the request or response. Another field differentiates electronic requests (97% of the total) from manual requests filed on paper and entered electronically by officials. While the publicly available data protect the identities of requesters by omitting names, user IDs, or answers to voluntary demographic questions, these data do include geographic information on a requester’s state and municipality, selected from “drop-down” menus of predefined options, thereby preventing misspellings.

While INFOMEX cannot prevent individuals from entering false information, misrepresentations of requesters’ locations appears to be relatively rare, as there is a strong relationship between municipality size and request volume. Across municipalities, the logged total number of requests filed and logged population are correlated at 0.76. If a large share of requesters were selecting false locations, we would not observe such a strong relationship. Further, any measurement error in this regard simply sets a higher bar for us to identify meaningful results.

The text of each request corresponds to the requester’s own open-ended description of the information they seek. While the majority of requesters described the nature of their requests within one of two primary text fields, roughly 13% of cases included either some or all of the request as a file attachment. We implemented an automated routine to download these based on links contained in the INFOMEX data, convert them to machine-readable text, and incorporate them into our corpus. We also removed all official requests for confidential requester personal information.<sup>6</sup> Finally, as a small number of requests were extremely long, often where an attached file contained an example report or a large spreadsheet, we truncated any request over 1,000 words, affecting only 0.02% of requests.

We applied a set of preprocessing techniques to convert our final corpus of requests into a collection of individual

5. This excludes state-level authorities and constitutionally autonomous bodies, as these fall under different legal regimes.

6. While these use INFOMEX, they are governed by different legal requirements, and additionally do not make public their request texts.

“documents” compatible with the unsupervised topic models used below, as well as to fix common typos, standardize the use of accents, and remove any nongraphical characters. We then processed each request text to remove all punctuation, numbers, names of months, and stopwords. The removal of these character sets is standard preprocessing for our topic model techniques (e.g., Bagozzi 2015; Roberts et al. 2014) and helps to eliminate extremely common, but highly uninformative, character strings. We also removed common phrases used by requesters to indicate that the information request is not provided in the text field, but rather is attached. Next, we omitted sparse terms that do not occur in at least 0.1% of the documents in our corpus, which is a common means of eliminating extremely rare words that offer little insight into the shared topics across one’s text corpus (Grimmer 2012; Quinn et al. 2010). Finally, all remaining words were converted to lower case, stemmed, and restructured into a document-term-corpus before analysis. Altogether, these preprocessing steps created a corpus with 1,023,267 processed request documents and 6,696 total unique words.

To uncover the different topics of interest to citizens making information requests, we use latent Dirichlet allocation (LDA; Blei, Ng, and Jordan 2003). LDA is an unsupervised topic model that identifies latent dimensions (topics) of texts. LDA estimates such topics via a hierarchical Bayesian model that treats documents as a finite mixture of underlying topics, and topics as a mixture over a latent set of topic probabilities (Blei et al. 2003, 993). In this manner, each identified topic can be characterized by a set of words that LDA determines as having the highest probability of association with that topic. These topic-word associations can then be used to (i) interpret each topic for its substantive meaning and/or (ii) to classify one’s documents in terms of the topic(s) most likely to have generated each document, based upon a document’s observed distribution of words.

With LDA, the analyst must explicitly choose the number of topics,  $k$ , to be estimated. In appendix B (apps. A–C are available online), we evaluate different approaches to this choice, and identify a  $k = 20$  as most appropriate for this application based on both qualitative evaluation and quantitative model fit statistics. We then estimate an LDA model on our full corpus of 1,023,267 preprocessed request texts. For this LDA model, in addition to setting  $k = 20$ , we assign the model’s  $\alpha$  parameter to favor the assignment of a single topic to each document,<sup>7</sup> and use Gibbs sampling for estimation. As such, this LDA model identifies the 20 top-

ics that best characterize the request texts across the entire 2003–15 period. We then classify each request according to its single most “dominant” associated topic (although we also utilize each request’s full set of twenty topic probabilities in our analyses).

In appendix C, we discuss an extensive set of evaluations aimed at interpreting these topics, including examination of the words most representative of each topic, and a close reading of several hundred individual requests sampled from those strongly associated with each topic. These interpretations are developed further in Berliner, Bagozzi, and Palmer-Rubin (2018). Based on these evaluations, we split these twenty topics into two sets: thirteen topics with higher public potential, and seven topics oriented toward more private goals. These are listed in table 2. We describe our LDA estimation strategy, and dominant topic approach, in further detail in appendices B and C.

Table 2. Twenty Topics of Requested Information

Topic Label	Percentage of Requests
<i>Environment and land</i>	8.64
Individual needs	8.50
<i>Budgets and spending</i>	7.98
Commercial information	7.86
<i>Distributive programs</i>	7.06
<i>Military, police, and crime</i>	6.84
Government employees 1: salaries and benefits	6.71
<i>Government employees 2: functions and qualifications</i>	6.26
<i>Energy and public utilities</i>	5.65
Internal procedures and official documents	5.19
Medical statistics	4.67
<i>Government employees 3: specific personnel</i>	4.10
<i>Public procurement 1: service providers</i>	4.09
<i>Banking, finance, and taxes</i>	3.83
<i>Education</i>	3.66
<i>Medical supplies 1: contracts and suppliers</i>	2.82
<i>Public procurement 2: procedures and documents</i>	2.32
<i>Public procurement 3: anticorruption</i>	2.03
<i>Medical supplies 2: purchases and spending</i>	1.24
Medical supplies 3: inventories	.56

Note. Twenty topics in order from largest to smallest after assigning each request to one primary topic, with the percentage of all requests. Topics designated as having higher public potential are set in italics.

7. We further justify this choice in app. B.

By topics with higher public potential we mean those that pose greater political risk to government entities. These include issues such as budgets and spending, crime and violence, environmental impacts, oil and utilities, distributive social programs, public procurement, and a topic capturing requests asking for information about specific public servants. These have clear politically relevant uses for investigation, oversight, media scrutiny, or political organizations. Requests on budgets and spending or on distributive programs might be aimed at identifying misuse of resources or favoritism. Requests about local environmental impact, permits for construction projects, or on decision procedures and compliance evaluations for government contracts might be aimed at uncovering corruption in allocating contracts and permits, or noncompliance with legal rules. Details on the scale of violence and the conduct of police and military in the “Drug War,” or on the qualifications of public officials, both also have clear potential to generate negative media attention or potential scandals. While even in these topics we still find some requests that are not likely filed with public goals, it is nonetheless these categories of requests that are most likely to activate a consideration of potential negative political consequences by officials.

Topics with lower public potential primarily reflect individual citizen needs, commercial interests, and demand for aggregate statistics. These pose less risk of generating negative media or investigative attention. For example, one topic largely comprises requests for information to fulfill individual citizen needs, such as how to access specific government services or how to complete specific procedures. Another primarily comprises requests for aggregate statistics pertaining to specific medical conditions. Relatively legalistic and sophisticated requests for commercial information, likely filed by companies themselves for private purposes, dominate a third topic. And while we identify other topics focusing on public employees as having higher public potential (due to a focus on specific individuals or qualifications that might be used to identify unqualified patronage hiring), another topic is primarily made up of requests seeking information on aggregate salary scales, benefits, and pensions—likely filed by employees themselves.

In appendix C, we describe all twenty topics in more detail, provide example requests, and explain our categorization of them as mainly “public” or mainly “private.” This division of topics is not absolute: an individual request for budget information might be mundane and apolitical, while a request pertaining to aggregate statistics might be aimed at evaluating government performance. However, these divisions nonetheless enable the testing of our secondary hypotheses by capturing the set of requests that are more likely to trigger political risk-management processes on the part of government officials. As a robustness check (app. A), we also analyze only the most

clearly private topic—Individual Needs—and the most clearly sensitive topic—Military, Police, and Crime—individually.

## MODELING GOVERNMENT RESPONSIVENESS

### Dependent variable

We are interested in government responsiveness to citizen information requests. In the context of information requests, responsiveness is the extent to which citizens receive the information they seek, in useful form, and in a timely fashion, except in cases legitimately subject to legal exemptions from disclosure. Given the complexities in defining and measuring these characteristics, we use information on both the timing and types of responses to triangulate across different possible measures of “nonresponsiveness.”

First, we measure the time to response in working days. In calculating this variable, we exclude weekends as well as official Mexican government holidays. However, given the presence of legal time limits for responses, we also employ two dichotomous measures of whether or not the response was late. While the statutory deadline was 20 working days, officials could request extension of up to 40 working days. We use both of these thresholds to construct separate indicators of late responses. Across the whole data set, 66.4% of requests received responses within 20 working days, while 89.3% of requests received responses within 40 working days. Although a late response may still provide the requested information, many information requests concern time-sensitive matters, particularly when publicly relevant. Further, delays often reflect a sensitive request that required greater intra-agency negotiation, either with the liaison unit or information committee.

We then employ three different measures of the type of response, based on information provided by officials’ categorization of the nature of their responses. Here we must confront two issues: First, that requests may be denied for legally allowable reasons, and second, that state agents themselves categorize their response types, and may not always do so accurately. Fox, Haight, and Palmer-Rubin (2011), for example, find that some responses (from the 2003–5 period) were categorized as positive with information delivered electronically, but the actual electronic attachments contain written justifications for not delivering the requested information. Nonetheless, they still concluded that for responses classified as positive, “information was actually provided in a satisfactory manner approximately 87 percent of the time” (Fox et al. 2011, 9).

We thus focus first on responses categorized as “Inexistencia de la información solicitada,” claiming that the requested information does not exist, making up 5.9% of all responses. According to Fox et al. (2011, 11), this is “the category of denial that leaves the greatest opening for agency opacity.” They found that many agencies denied requests under this category

without written justification, concluding that “agencies have discovered that this is the least risky way to deny requests for information in cases where they did not want it released or when assembling it would be a large burden,” especially since “in contrast, the burden of proof is on the agency if it claims that information requested is ‘confidential’ or ‘reserved’” (Fox et al. 2011, 14). Thus, while a denial for other reasons may well be legitimate under the law, we consider the “Inexistencia” responses to more clearly be examples of unresponsive state agents.

Second, we construct another dependent variable for responses categorized as either “Negativa por ser reservada o confidencial” or “Información parcialmente reservada o confidencial,” indicating that the information requested is either fully or partially classified or confidential. Although many of these denials will be justified, there may be political bias on the margins. These responses account for 3.6% of the total.

Finally, we construct a dependent variable to capture situations where officials do disclose the requested information, but in a nonelectronic form that creates additional financial or physical barriers to the requester. Whereas the majority of positive responses are categorized as “Entrega de información en medio electrónico,” meaning the information was delivered electronically, some are physically shipped (requiring the requester to first pay a fee), while others require the requester to physically appear at a government office. These place additional burdens on the requester (and further compromise their anonymity) that in some cases may be prohibitive—indeed several thousand requests appear abandoned in the INFOMEX system, without the requester having ever paid the shipping fee. Lagunes and Pocasangre (2019, 169), discussing the burdensome procedure required to make such payments, suggest that “fee charges appear to be a strategy by which to discourage citizens from pursuing a request.” We capture these cases with three categories of responses: “Notificación de envío” (shipping notification), “Notificación lugar y fecha de entrega” (place and date to pick up the information), and “Respuesta del solicitante a la notificación de entrega de información con costo” (awaiting the requester paying the fee). Together, these types of responses account for 4.9% of the total.

Each of these dependent variables captures different dimensions of responsiveness on which we might expect to see political bias in accordance with our hypotheses. As an initial visualization of these measures of responsiveness in relation to political alignment, figures 1 and 2 plot each measure against state-level vote shares.<sup>8</sup> Figure 1 plots average re-

sponsiveness to requests in each state, over the period up until the 2012 presidential election, against the state-level PAN vote share averaged between the 2000 and 2006 presidential elections. Figure 2, in turn, plots average responsiveness over the period after newly elected PRI President Enrique Peña Nieto took office in December 2012, against state-level PRI vote share in the 2012 presidential election. In each case except one, the line-of-best-fit slopes downward, suggesting that areas of higher governing party support tended to receive better responses. However, our subsequent approach moves beyond the state level to use the individual request as the unit of analysis.

### Independent variables

Our primary independent variable captures political alignment at the municipality level. For each information request, this variable takes the value of the governing party’s vote share in the municipality of origin, in the previous presidential election. Here, we first match each request to the municipality from which it was sent and to corresponding municipality-level data on presidential election vote shares. The resulting variable, Government Vote Share, reflects PAN vote share from the elections of Vicente Fox (2000–6) and Felipe Calderón (2006–12), and PRI vote share from the election of Enrique Peña Nieto (2012–18).

We control for the logged population of each municipality to avoid confounding by differences in requests from large urban areas and smaller rural locales. We also control for the logged “economic marginality” of each municipality, an index of economic and demographic indicators, to ensure that the effect of governing party political support is not merely proxying for underlying economic differences.

We also include several request-level measures, in order to capture features of requests that may influence the nature of official responses. First, we control for all twenty sets of topic probabilities, reflecting the extent to which each request contains words associated with each of the twenty different potential topics of request texts. We include indicators of whether or not each request included an attached file, and the medium of each request (electronic or physical), although nearly all are electronic. We also measure the (logged) word count of each request, and the (logged) ratio of characters to words. This ratio of characters to words is included as a proxy for “readability,” a measure of the complexity of text.

We also measure two other features of requests that might reflect their complexity or sophistication—legalism and punctuation.

8. We omit Morelos, as a Morelos-based group filed thousands of highly unusual requests (see app. C) whose more negative responses heavily influ-

enced the state’s averages. While we do not omit Morelos from our subsequent analyses, our results are similar if we do.

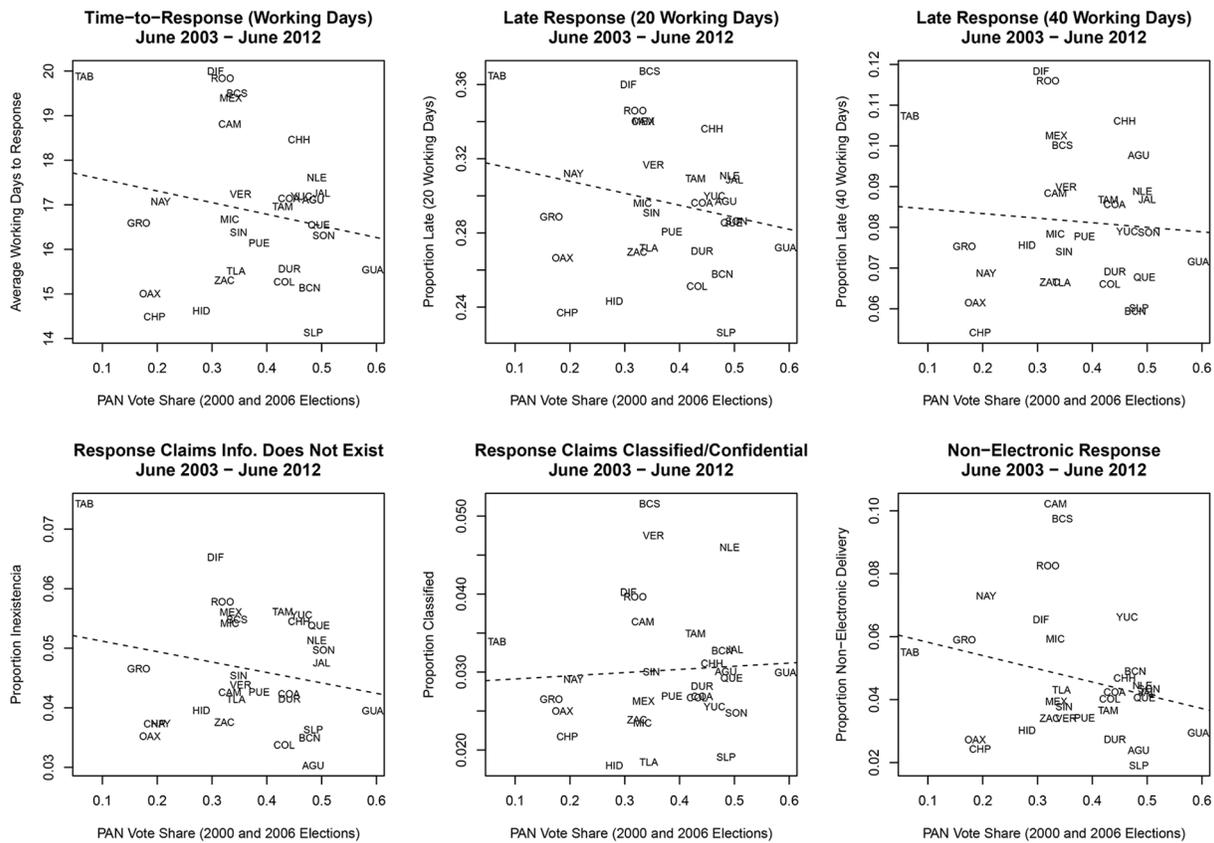


Figure 1. State-level averages for six different measures of responsiveness, plotted against state-level governing party vote shares. Response measures are averages for all requests filed in each state over the period up until the 2012 presidential election. Vote shares are PAN vote share averaged between the 2000 and 2006 presidential elections. Dashed line is a line of best fit from a linear regression of response (state averages) on state vote share. Morelos omitted as an outlier (see n. 8).

Some information requests are highly legalistic, citing provisions of the ATI law or otherwise conveying a level of expertise on the part of the requester that might influence the ways that officials respond. Other requests are clearly filed by individuals with little political sophistication, use very ordinary language, and often omitting punctuation. To capture legalism, we first count the number of times each request includes a series of words or phrases that reflect references to the ATI law or related legal provisions,<sup>9</sup> and then measure this count as a ratio of the length of each request. We also measure the incidence of punctuation relative to the length of each request.<sup>10</sup>

To measure the investigative risk of individual requests in a more fine-grained manner beyond topics, we constructed a “Corruption Discourse Index” using a dictionary of words drawn from newspaper articles. We first identified over 300,000 articles from two major newspapers (*Reforma* and *El Universal*)

mentioning any of 20 large federal agencies with unambiguous names and abbreviations. We then developed a set of 18 search terms pertaining to corruption or scandal, and identified the set of articles containing such terms. We used these to produce a dictionary of 4,568 words that were distinctive to the corruption articles, as compared to the noncorruption articles about the same agencies. Each information request, in turn, was scored with the proportion of words that matched this dictionary. As expected, this measure on its own is strongly associated with worse responses.

We also control for a potentially relevant time-varying factor at the agency level, calculating a measure of “agency workload” to capture the (logged) number of requests received over the preceding 30 days. For each agency, this variable takes a different value on each day of the period under study, reflecting the fact that officials facing a larger “queue” of pending requests may respond in different ways. We also include indicators for each month of the year, to capture any seasonal trends in response patterns.

Finally, we include a set of fixed effects for each government agency (thereby making only within-agency comparisons for requests from different regions of the country), each

9. These are: “ley federal de transparencia,” “ley de transparencia,” “artículo,” “fracción,” “sujeto obligado,” and “LFTAIPG”—both with and without accents where relevant and regardless of case.

10. The punctuation we consider is as follows: periods, commas, colons, semicolons, question marks, and inverted question marks.

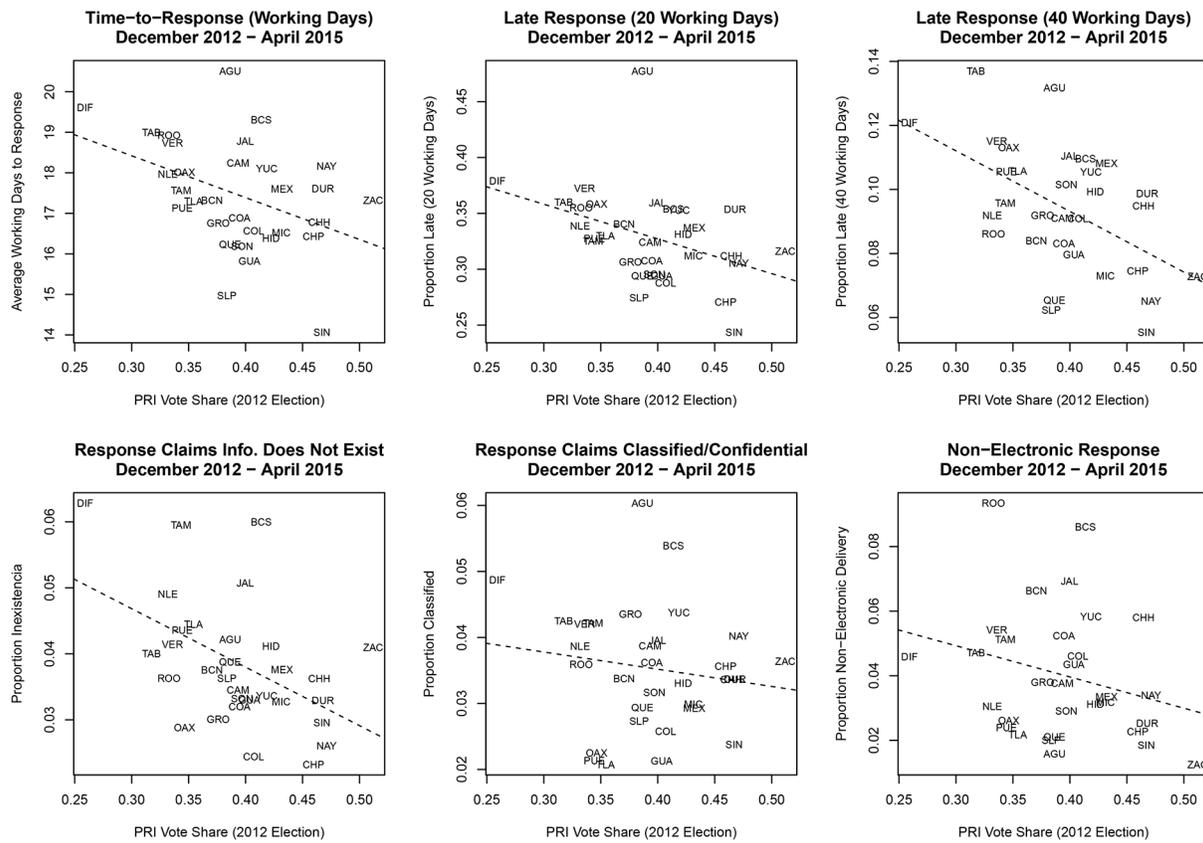


Figure 2. State-level averages for six different measures of responsiveness, plotted against state-level governing party vote shares. Response measures are averages for all requests filed in each state over the period after newly elected PRI president Enrique Peña Nieto took office in December 2012. Vote shares are PRI vote share in the 2012 presidential election. Dashed line is a line of best fit from a linear regression of response (state averages) on state vote share. Morelos omitted as an outlier (see n. 8).

year (to capture cross-the-board time trends), and each “dominant” topic category (reflecting the topic with the highest probability of association with each request). Rather than including these fixed effects separately (which we show in app. A), we employ a separate fixed effect for every single agency-topic-year combination. That is, our models make comparisons only among requests filed with the same agency, on the same topic, and in the same year—but from different parts of the country with varying political alignment. For example, comparisons are made among requests filed with the environmental ministry, on the topic of budgets and spending, in the year 2008—but not across different topics of information in requests filed with that same ministry, and not across requests for budget and spending information in requests filed with different ministries, and not across different time periods. This distinction is important, as we do not wish to assume that each government agency has the same baseline procedures for responding to different topics of requests, or to assume that changes in internal procedures over time took place according to identical time lines across agencies. These very fine-grained fixed effects automatically capture the effect of any omitted variable that varies across agency-topic-year combinations, and allow us to

more closely approximate the level of comparability that would be obtained in an experimental approach.

### Sample and modeling approaches

Rather than using the entire sample of over 1 million requests, we carefully select the set of requests for which our hypotheses are most appropriate. First, we omit the roughly half of all requests originating from the Federal District, where Mexico City is located. Given the preponderance of media, political, legal, and civil society organizations filing requests in the capital, we would be concerned that these requests are less comparable with those in the rest of the country, would make the municipality of origin a less reliable proxy for the likely political alignment of the requester, and would potentially lead to misleading results given the high support for the PRD in the capital city. However, results remain highly similar when we instead retain Federal District observations (app. A). Second, we omit agencies that only rarely receive requests, as their responses are likely to be more idiosyncratic. We include only agencies that received, on average, more than one request per week. While this retains only 189 of 304 possible agencies, it removes only 36,074 requests, a small fraction of the total.

Results remain highly similar when we retain these (app. A). Third, to avoid any bias from “right-censoring,” we remove any requests filed after May 1, 2015, or otherwise unresponded to by our August 2015 cutoff. Finally, we omit all requests filed in the period between July 1, 2012—the date of the 2012 presidential election—and the end of 2012 when the new PRI President Enrique Peña Nieto took office.<sup>11</sup> For purposes of evaluating the effects of political alignment, this period is likely to be idiosyncratic, and it is not clear which party’s supporters we would expect potential political bias to favor. The resulting sample includes 456,747 individual requests.

We employ two main modeling approaches. The first is to model responses as a function of the previously mentioned covariates, making cross-municipality comparisons between requests sent from areas of varying political alignment. By including fixed effects for each agency-topic-year combination, we account for potential unobserved confounders on those dimensions. Our second approach instead leverages changes in political alignment over time, adding fixed effects for municipalities and removing fixed effects for each year. This approach takes into account only the over-time variation in political alignment between municipalities and the central government, primarily arising from the change in the governing party after 2012. For both these approaches we employ linear probability models, as employing categorical models with such a large number of fixed effects poses problems for both computation and interpretation.<sup>12</sup>

## RESULTS

Table 3 presents our main set of results, including the previously discussed control variables and fixed effects for every agency-topic-year combination. For five of the six dependent variables, the coefficient for government vote share is negative and statistically significant at  $p < .05$ . Only for classified/confidential responses is the coefficient not significant. Notably, the inclusion of fixed effects means that this analysis already accounts for any unobserved variation at the level of agencies, topics, years, or any combinations thereof, as well as controls for other municipality-level factors, request-level complexity, sensitivity, and sophistication, and the time-varying workload of each agency. These results offer strong support in favor of the Partisan Discretion hypothesis. Ministry personnel do in fact appear to exercise discretion in responding to information requests, favoring request from

areas of core support. But which political logic drives this discretionary behavior—favoring allies with access to benefits, or keeping potentially damaging information out of the hands of opponents?

To adjudicate between the Targeting Benefits hypothesis and the Mitigating Risk hypothesis, we examine whether political effects of government vote share appear differentially for topics of requests with higher and lower public potential. Table 4 presents results of models restricting the sample to only topics we categorized as having higher public potential (this and all subsequent tables display only the main coefficients of interest, although all controls from table 3 are still included). All negative coefficients are greater in magnitude. Table 5 restricts the sample to only topics categorized as having lower public potential. In this case, only one of the coefficients is statistically significant.<sup>13</sup> These results suggest that greater political targeting of responses takes place for topics of information that pose potential political risks, and not for politically mundane or purely private topics of information. This finding is thus more consistent with the Mitigating Risk hypothesis than the Targeting Benefits hypothesis. In appendix A, we also show results employing an interaction term between Government Vote Share and the indicator of higher public potential. The coefficient on the interaction term is negative in all cases, and statistically significant for four of the six models.

In appendix A, we present additional robustness checks for our main findings from table 4. First, we include in the sample the requests from the Federal District, which do not substantially change our results. Second, we measure the main independent variable in several alternate ways, including using state-level rather than municipal-level vote shares (in case officials incorporate information about political alignment at a more geographically aggregated level), or measuring the governing party’s vote margin over the next-largest party rather than vote share, or using vote shares from mayoral elections in each municipality (though these are more limited in data availability) rather than in presidential elections. Results from these models remain largely in line with our primary results. Third, we control for measures that capture the extent to which a locale represents a “swing” district, as a potential alternate hypothesis to ours might be that officials target better responses to such areas, in an attempt to provide valuable services or better demonstrate government performance. We alternately use the Effective Number of Parties in presidential vote shares of each municipality, or the closeness

11. Although the new president took office December 1, 2012, we exclude requests through December 31, 2012, to ensure that indicators for 2012 do not extend across two separate administrations.

12. In app. A, we show logit results for robustness. In all models, we employ two-way clustered standard errors by both municipality and agency.

13. A coefficient equality test shows that the coefficients for government vote share in tables 4 and 5 are significantly different from each other for models 1, 2, and 6.

Table 3. Full Models of Government Responsiveness to Information Requests

Dependent Variable	Model 1 LogDays	Model 2 Late20	Model 3 Late40	Model 4 Inexistencia	Model 5 Classified	Model 6 Nonelectronic
Government vote share	-.124** (.038)	-.030* (.013)	-.022** (.008)	-.017** (.006)	.004 (.004)	-.012* (.005)
Log population	.007 (.006)	-.001 (.002)	-.005 (.003)	-.002 (.002)	-.000 (.001)	.002* (.001)
Log econ. marginality	-.049** (.015)	-.016** (.005)	-.016** (.004)	-.005* (.002)	-.002 (.001)	-.004 (.003)
Request length	.189** (.019)	.065** (.007)	.032** (.003)	.010** (.001)	.007** (.002)	.010** (.002)
Request readability	.106** (.032)	.045** (.014)	.030** (.007)	-.006 (.005)	.015* (.007)	.033** (.010)
Request with attachment	-.169** (.018)	-.043** (.007)	-.017** (.006)	-.007 (.005)	-.011** (.004)	-.003 (.005)
Request medium	-.038 (.054)	-.071** (.021)	-.023** (.008)	-.013* (.005)	.011 (.009)	.070 (.038)
Request legalism	1.060 (.656)	.521 (.330)	.219 (.184)	.369 (.259)	.183 (.179)	-.060 (.211)
Request punctuation	.365** (.089)	.060 (.033)	-.009 (.026)	-.008 (.016)	.001 (.016)	.028 (.016)
Request corruption discourse	.154* (.069)	.097** (.026)	.055** (.010)	.014 (.013)	.048** (.012)	.002 (.009)
Log agency workload	.057** (.012)	.030** (.005)	.014** (.005)	.008* (.004)	-.003 (.003)	-.008** (.002)
Dependent variable mean	2.561	.332	.100	.056	.032	.042
Observations	456,747	456,747	456,747	456,747	456,747	456,747
Adjusted R <sup>2</sup>	.323	.240	.161	.123	.092	.121

Note. Models additionally include 20 topic probabilities, monthly indicators, and fixed effects for every agency. Two-way clustered standard errors on municipality and agency.

\*  $p < .05$ .

\*\*  $p < .01$ .

Table 4. Models of Government Responsiveness to Information Requests, with Sample Restricted to Topics with Higher Public Potential

Dependent Variable	Model 1 LogDays	Model 2 Late20	Model 3 Late40	Model 4 Inexistencia	Model 5 Classified	Model 6 Nonelectronic
Government vote share	-.162** (.040)	-.049** (.014)	-.030** (.009)	-.019** (.007)	-.000 (.005)	-.021** (.004)
Dependent variable mean	2.593	.346	.109	.062	.033	.038
Observations	305,115	305,115	305,115	305,115	305,115	305,115
Adjusted R <sup>2</sup>	.334	.251	.173	.121	.096	.098

Note. Models additionally include the same control variables as in table 3, 20 topic probabilities, monthly indicators, and fixed effects for every agency-topic-year combination. Two-way clustered standard errors on municipality and agency.

\*  $p < .05$ .

\*\*  $p < .01$ .

Table 5. Models of Government Responsiveness to Information Requests, with Sample Restricted to Topics with Lower Public Potential

Dependent Variable	Model 1 LogDays	Model 2 Late20	Model 3 Late40	Model 4 Inexistencia	Model 5 Classified	Model 6 Nonelectronic
Government vote share	-.052 (.053)	.005 (.019)	-.009 (.012)	-.014* (.006)	.012 (.007)	.006 (.011)
Dependent variable mean	2.500	.303	.083	.045	.030	.050
Observations	151,632	151,632	151,632	151,632	151,632	151,632
Adjusted R <sup>2</sup>	.300	.213	.129	.127	.088	.158

Note. Models additionally include the same control variables as in table 3, 20 topic probabilities, monthly indicators, and fixed effects for every agency-topic-year combination. Two-way clustered standard errors on municipality and agency.

\*  $p < .05$ .

\*\*  $p < .01$ .

of the vote margin between the nationally governing party and the largest other party. In either case, the coefficient for the “swing” variable is close to zero, and not statistically significant, in all six models.

Fourth, we use alternative topic models, setting the number of topics to 19 or 21, instead of 20, to address potential concerns over topic instability (Wilkerson and Casas 2017). Fifth, we employ separate fixed effects for each agency, topic, and year (rather than for every combination thereof). Sixth, we instead use more fine-grained fixed effects that make comparisons only among requests with identical agency, topic, year and similar “level of difficulty.” Seventh, we take into account potential relationships among the different outcome variables by controlling for measures of response type in our models of response time, and vice versa. Finally, we include additional request text-based control variables, assessing the prevalence of misspellings and irregular capitalization. In all of these checks, the results remain substantially similar.

We next turn to a different approach, employing municipality fixed effects to account for any unobserved municipal-level characteristics, while also identifying results based solely on over-time changes within each municipality. That is, the coefficient on government vote share in these models is primarily informed by changes at the 2012 election, before which its value reflects the PAN vote share and after which its value reflects the PRI vote share in each municipality. If our first hypothesis is correct, then requests from a largely pro-PAN municipality should receive more favorable responses before 2012, and less favorable responses afterward, and vice versa for a largely pro-PRI municipality. Table 6 presents the results of models including municipality fixed effects, in addition to the other control variables in the models from table 3, as well as fixed effects for every agency-topic combination. We do not incorporate fixed effects for year, in order to capture within-municipality variation over time in political alignment. In four of the six models, the coefficient on government vote

Table 6. Models of Government Responsiveness to Information Requests, including Municipality Fixed Effects

	Model 1	Model 2	Model 3	Model 4	Model 5	Model 6
Government vote share	-.291** (.058)	-.143** (.030)	-.065** (.020)	-.035** (.011)	.001 (.007)	-.014 (.010)
Dependent variable	LogDays	Late20	Late40	Inexist.	Classif.	NonElec.
Dependent variable mean	2.561	.332	.100	.056	.032	.042
Observations	456,747	456,747	456,747	456,747	456,747	456,747
Adjusted R <sup>2</sup>	.263	.194	.130	.086	.067	.099

Note. Models additionally include the same control variables as in table 3, 20 topic probabilities, monthly indicators, and fixed effects for every agency-topic combination. Two-way clustered standard errors on municipality and agency.

\*  $p < .05$ .

\*\*  $p < .01$ .

share is statistically significant, in all four cases at a  $p < .01$  level. We further test our subsidiary hypotheses in the municipality-fixed effects approach (app. A), and find that the effects of political alignment are stronger among topics of information with higher public potential, than among those with lower public potential.

To assess the substantive sizes of these effects, we compare the responses to hypothetical requests from municipalities with governing party vote shares of 0.25 and 0.75, based on the main results in table 3. A request from the more government-opposed municipality would be 4.9% more likely to receive a response later than in 20 working days, 13.5% more likely to receive a response later than in 40 working days, 19.8% more likely to receive a response claiming that the information does not exist, and 12.8% more likely to receive a nonelectronic response.<sup>14</sup> Further, these effects are larger for the sample of topics with higher public potential.

We emphasize that these represent average effects, across a wide variety of agencies, time periods, and topics of information. Many requests will be routine, and pose no chance of activating a consideration of political risk on the part of officials. Only for some subset of requests would we expect any chance of officials considering the identity of the requester or the political alignment of their location in the first place. We are thus capturing differences averaged across a set of requests that falls within the scope conditions of our argument, and a set of requests that falls outside those scope conditions. While we differentiate between two sets of requests based on the likely public potential of their topics, this can only capture a partial level of detail.

Mindful not to engage in “cherry-picking” subsets of our data without a priori theoretical reasons, we do offer one illustrative example of a subset of information requests with results indicating much stronger effects of local political alignment. In appendix A, we present results for a sample of only requests filed with *Petróleos Mexicanos*, or PEMEX, the state oil company (and again omitting requests filed from the Federal District). All coefficients are much larger in magnitude than in the main results from table 2. Comparing hypothetical requests from municipalities with varying vote shares as above, our results suggest that a request from the more government-opposed municipality would be 19% more likely to receive a response later than in 20 working days, 31.1% more likely to receive a response later than in 40 working days, 44.6% more likely to receive a response claiming that the infor-

mation does not exist, and 78.6% more likely to receive a nonelectronic response. This illustration suggests that the extent of political discrimination is much stronger than our primary results suggest for some agencies, types of information, or specific time periods. On the other hand, many other types of requests are routine and likely to be handled in a purely impartial manner, and some entire agencies may be fully professionalized and impartial such that they allow no discrimination whatsoever.

## CONCLUSION

We study responsiveness in the provision of government information, using publicly available data containing every information request filed with Mexican federal government agencies from 2003 to 2015. Similar to studies of distributive politics, we analyze how responsiveness is shaped by discretionary partisan bias. Yet while something of a consensus in the distributive politics literature has emerged that politicians target core supporters with private benefits—useful in clientelistic exchange—more than with public goods, we find the opposite. Officials favor areas of core support with information that lends itself to generating public accountability and demonstrate little to no partisan bias in the provision of information useful in generating private benefits for the requester.

Underlying these findings are multiple countervailing incentives confronting the agents of accountability who are charged with responding to information requests. Although ostensibly institutions of downward accountability, as well as horizontally accountable to the oversight body IFAI, officials are also accountable upward to their agency heads and politicians with electoral incentives. These multiple forms of accountability create different motivations for and against disclosure in different circumstances. Favoring disclosure, they may be motivated by an interest in sustaining a positive image of their ministries, career incentives based on performance, sanctions for noncompliance, or personal commitment to transparency. The main incentives disfavoring disclosure are the political risk associated with releasing evidence of corruption or poor governance to the general public, along with the needs to comply with official exemptions, and cope with limited capacity. While agencies certainly could exercise partisan discretion in delivering information useful for reaping private benefits, they instead exercise caution in the provision of information that has public potential, given the high political cost associated with scandals.

By incorporating information on the topics, complexity, sensitivity, and sophistication of requests, and making comparisons only among requests filed with the same government

14. Percentage increases on the baseline probability of each outcome. We skip classified responses as we find no significant effects in line with our expectations for this outcome type.

agency, on the same topic, and in the same year, we address concerns over underlying differences in the requests themselves. We find evidence that municipalities with higher vote shares for the president's party receive more favorable responses, across multiple measures of responsiveness. However, this finding holds only for requests on topics of information with higher public potential, and not for more private topics of information.

Our findings demonstrate that, even in a relatively routine mode of citizen-government interaction featuring decentralized decisions made by large numbers of officials, ruling-party interests and the need to manage political risks can generate dynamics of responsiveness favoring areas of core government support. These findings also suggest that, despite formal rules intended to preserve requester anonymity and procedural fairness in Mexico, informal political dynamics shape bureaucratic processes and allow political discrimination to take place. Yet, on the other hand, we find no evidence of a clientelistic logic in the discretionary disclosure of information to confer private benefits. This suggests some success in insulating accountability mechanisms from patronage politics, yet their replacement with partisan bias of a different form.

Our results, we believe, are applicable well beyond the case of Mexico. We would expect to find similar bias in responsiveness to citizen information requests in any system featuring the means (functioning ATI institutions), motive (electoral or institutional sanctions for politicians), and opportunity (bureaucratic discretion and compromised anonymity) for such discrimination. This includes a wide variety of democratic and nondemocratic regimes. To be sure, geographic targeting on the basis of electoral returns is not the only strategy available to officials to discriminate in disclosure; they may also base decisions on other cues that signal the provenance of the request, such as the sophistication of the language, signals of insider knowledge, or even the precise identity of the requester—as in Roberts's (2005) study of Canada. However, given the politically damaging nature of corruption revelations in even the most mature democracies, we expect the filtering of potentially damaging information to be a primary consideration in discretionary disclosure.

Reforms aimed at minimizing such abuses could include measures to insulate bureaucrats charged with handling information requests from partisan motives, strengthen oversight agencies to detect and sanction noncompliance, and better procedural protections for citizen anonymity. However, such policies are typically much easier to propose than to meaningfully establish.

More broadly, the problem of discretionary responsiveness in institutions of accountability is a particularly daunting challenge for all democracies, transitional democracies in par-

ticular. Institutions created or strengthened through “second-generation” democratization—such as anticorruption agencies, state auditors, and judiciaries—have proven quite resistant to attempts to augment their autonomy and aggressiveness (O'Donnell 1998). Our findings extend this concern to transparency institutions, demonstrating how an institution designed to empower citizens to hold politicians accountable can be utilized by state actors as a tool of selective responsiveness.

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